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Published in:
IEEE Transactions on Audio, Speech and Language Processing

DOI (link to publication from Publisher):
10.1109/TASLP.2014.2367817

Publication date:
2015

Document Version
Publisher's PDF, also known as Version of record

Link to publication from Aalborg University

Citation for published version (APA):

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Joint Pitch and DOA Estimation Using the ESPRIT Method

Yuntao Wu, Member, IEEE, Leshem Amir, Senior Member, IEEE, Jesper Rindom Jensen, Member, IEEE, and Guisheng Liao

Abstract—In this paper, the problem of joint multi-pitch and direction-of-arrival (DOA) estimation for multichannel harmonic sinusoidal signals is considered. A spatio-temporal matrix signal model for a uniform linear array is defined, and then the ESPRIT method based on subspace techniques that exploits the invariance property in the time domain is first used to estimate the multi-pitch frequencies of multiple harmonic signals. Followed by the estimated pitch frequencies, the DOA estimations based on the ESPRIT method are also presented by using the shift invariance structure in the spatial domain. Compared to the existing state-of-the-art algorithms, the proposed method based on ESPRIT without 2-D searching is computationally more efficient but performs similarly. An asymptotic performance analysis of the DOA and pitch estimation of the proposed method are also presented. Finally, the effectiveness of the proposed method is illustrated on a synthetic signal as well as real-life recorded data.

Index Terms—DOA estimation, ESPRIT, performance analysis, pitch estimation.

I. INTRODUCTION

The problem of estimating the pitch of a periodic signal is a fundamental problem that has received considerable attention for many years due to its wide application in areas such as audio and speech coding, compression, enhancement and classification of music, and speech analysis [1]–[5]. Many methods for parameter estimation in both single-pitch and multi-pitch scenarios have been reported (see [6]–[8] for an overview of existing pitch estimation methods). Direction-of-arrival (DOA) estimation with an array of spatially separated sensors is another key research topic in the field of signal processing [10], [11], which has been widely studied in the past decades. Its application areas include radar [12], sonar [13], radio astronomy [14], geophysics [15], and speech processing with a microphone array [16]. In recent years, the problem of joint estimation of pitch and DOA of audio and speech signals received by multiple sensors has been extensively discussed for both single-pitch and multi-pitch scenarios in the literature [17]–[36]. Some of the key examples of applications that could benefit from this includes teleconferencing, surveillance applications, hands-free communication, and hearing aids, and so on.

The problem of single-pitch estimation has been researched thoroughly, but the multi-pitch scenario occurs more often in real-life applications. The performance of multi-pitch estimators designed for single channel recording is seriously degraded in the presence of frequency overlapping of sources, reverberation, and strong noise environments. For these reasons, the problem can be solved using a multi-channel receiver approach to obtain an accurate pitch estimation in such cases. Some robust multi-pitch estimation methods have been proposed by using joint spatio-temporal processing [17], [18], where joint parameter estimation is performed, and hence the sources are still resolvable even in the case of a similar source parameter in one dimension for both sources; i.e., with nearly same DOA or pitch. For example, if the pitch of two periodic sources, such as many audio and voiced speech, is similar, we can use the distinct spatial information to first estimate the DOA and then extract the pitch from the spatially separated sources [33]. Most of the methods for joint DOA and pitch estimation focus on the frequency and the DOA estimation of a single sinusoid defined in two dimensions, such as the space-time realization approach [31], the 2-D minimum variance distortionless response (MVDR) method [22], [23], [24], the 2-D ESPRIT method [40], and so on. Generally, these methods are based on the assumption of a narrow-band signal model, therefore, these methods are not suitable for the speech and audio processing since the problem of parameter estimation of audio and speech recorded using a microphone array is considered as a broadband case. Other methods can be considered as a generalization of the previous methods by assuming that desired sources are modelled as sums of harmonically related sinusoids, which is a good fit for many musical instruments and voiced speech. In other words, the traditional broadband case is transformed into multiple narrow-band cases. Some classic methods include the ML-based method [28], subspace-based methods [27], and the linearly constrained minimum variance (LCMV) beamformer.
[32]; however, these are focused on time-delay and pitch estimation, and the time-delay estimates are closely related to DOA estimation. To solve the problem of DOA and pitch estimation in the broadband signal case, one common form of processing is to use band-pass filtering of the broadband signals into multiple subbands. By doing so, the narrow-band estimators can be applied to each subband [20], [21], [40]. Since the harmonic signals consist of sinusoidal components, an alternative method is to model each source as a multiple narrow-band signal with distinct frequencies with the same DOA. In recent years, based on the harmonic model, some methods for joint pitch and DOA estimation for single-pitch and multi-pitch scenarios have been proposed. Although the maximum likelihood (ML) estimator [28] provides an optimum parameter estimation, its computational complexity is extremely demanding. A non-linear least squares (NLS) method for joint DOA and single-pitch estimation was recently reported in [36], which is also equivalent to the ML estimator under the assumption of white Gaussian noise and an anechoic environment. To avoid the multidimensional grid search requirement of both the ML and NLS-based methods, a simpler but suboptimal solution can be achieved by the subspace-based approach, which relies on the singular value decomposition or eigen-decomposition of the observation space into signal and noise subspaces. However, conventional subspace-based techniques, such as the multi-channel multi-pitch harmonic multiple signal classification (MC-HMUSIC) method for joint DOA and pitch estimation still requires a 2-D grid search [34]. In order to reduce the computational complexity of the two-dimensional searching used in the MC-HMUSIC method, an improved subspace-based technique known as multi-channel optimally weighted harmonic multiple signal classification (MCOW-HMUSIC) was recently proposed in [35], where the 2-D searching for each source in the MC-HMUSIC method is converted into two 1-D searching computations. Furthermore, many of the above-mentioned methods involving grid searches require reasonable initial values of the pitch or DOA estimates or a fine estimation grid to get an accurate estimation. This technique still suffers from the problem of finding the correct coupling between the DOA parameters and the pitch parameters. Furthermore, it is well known that the 2D MUSIC cannot easily cope with multi-path.

The aim of this paper is to develop a computationally efficient pitch and DOA estimation algorithm by utilizing the structure of a spatio-temporal signal model. The conventional ESPRIT method is exploited to estimate the pitch and DOA of interest for the single-pitch and multi-pitch scenarios, which is also a subspace scheme but does not involve any searching, and hence has a lower computational complexity than methods based on MC-HMUSIC [34] with only negligible performance loss. Moreover, an important contribution of this paper is the introduction of spatial smoothing techniques, that allow us to resolve multi-path propagation. This makes the proposed technique much more practical in real life environments as we demonstrate. In the presence of multi-path propagation, we propose to generalize the spatial smoothing technique [39] to spatio-temporal smoothing which facilitates the resolvability of both pitch frequencies in the multi-path case.

The rest of the paper is organized as follows. The problem formulation is firstly described in Section II. In Section III, the spatio-temporal data model for the single-pitch case is defined and the single-pitch and DOA estimation method based on the ESPRIT technique is given. Then, the generalization of the proposed method to the case of joint multi-pitch and DOA estimation is presented. This is followed by the development of a joint pitch and DOA estimation algorithm in the presence of multi-path propagation. Section IV summarizes the proposed methods, and the asymptotic performance analysis follows in the subsequent section. Finally, simulation results and real-world signal data tests are discussed to illustrate the effectiveness of the proposed method. Conclusions are drawn in Section VII.

Notation: The following notations are used in this paper: bold, lowercase letters refer to column vectors and uppercase bold letters refer to matrices. The inverse, pseudo-inversion, transpose, Hermitian transpose and conjugation of a matrix are denoted by $(\cdot)^{-1}$, $(\cdot)^{+}$, $(\cdot)^{T}$, $(\cdot)^{H}$ and $(\cdot)^{*}$, respectively. The statistical expectation operator is referred to as $\mathbb{E}\{\cdot\}$. The Kronecker product is denoted by $\otimes$ and the vectorized representation of a matrix is given by $\text{vec}(\cdot)$ respectively. The diagonal matrix consisting of only the diagonal entries of a square matrix is given by $\text{diag}\{\cdot\}$. $\text{Re}\{\cdot\}$ stands for the real part of a complex number and $\angle\{\cdot\}$ denotes the phase of $(\cdot)$, respectively. $\mathbb{I}$ is the identity matrix and $\mathbb{0}$ is a matrix of zeros. Estimated parameters are denoted by a hat, $\hat{\cdot}$.

II. Problem Formulation

To demonstrate the proposed solution without complicating the notation, we begin by demonstrating how the joint DOA and pitch can be estimated when the received signals are incoherent, i.e., there is no multi-path propagation. The multi-channel signal model is presented as follows. Assume that the signal $x_r(n)$ received by the $r$-th microphone element arranged in a uniform linear array (ULA) configuration with inter-element spacing $d$, for $r = 0, \cdots, R - 1$, is given by [34], [36]:

$$x_r(n) = s_r(n) + q_r(n), \quad n = 0, 1, \cdots, N - 1, \quad (1)$$

$$s_r(n) = \sum_{p=1}^{P} \sum_{i=1}^{L_p} \alpha_{p,i} e^{j(\omega_p n + \phi_{p,i})}, \quad (2)$$

where $N$ denotes the number of samples, and $P$ and $L_p$ are the number of spatial sources and the model order (harmonic number) of the $p$-th spatial harmonic source, respectively. Moreover, $\omega_p$ denotes the pitch of the $p$-th harmonic source, $\alpha_{p,i}$ is the complex-valued amplitude of the $l$-th harmonic of the $p$-th source and $q_r(n)$ stands for additive white Gaussian noise with variance $\sigma^2$ received at the $r$-th sensor. The phase shift caused by the propagation time-delay between array elements is defined by $\phi_{p,i} = \omega_p f_s \frac{c}{2} \sin(\theta_p)$, where $c$ is the speed of propagation, $f_s$ is the signal sample frequency, and $\theta_p$ is the DOA (direction-of-arrival) of the $p$-th spatial harmonic source. In this paper, the number of sources $P$ and the number of harmonics $L_p$ of each source $p$ are assumed to be known a priori, or determined with some available methods such as [6],
The objective of this paper is to estimate the nonlinear parameters \( (\omega_p, \phi_q), p = 1, 2, \ldots, P \) from the received data \( x_r(n) \). It is worth noting that parameter \( \phi_q \) is dependent on the fundamental frequency \( \omega_p \). Meanwhile, when the multi-path propagation is present, the model in (1)-(2) remains the same except that we allow identical pitch with \( \omega_p - \omega_q \) for \( p \neq q \).

III. PROPOSED METHOD

A. Spatial-Temporal Signal Model in the Single-Pitch Case

To present the proposed method clearly, we first focus on the problem of parameter estimation in the single-pitch case, that is where \( P = 1 \) in the signal model (1). In the signal model (1), for \( P = 1 \), the received array data can be written as

\[
x_r(n) = s_r(n) + q_r(n), n = 0, 1, \ldots, N - 1,
\]

\[
s_r(n) = \sum_{l=1}^{L} \alpha_l e^{j\omega_l n + \phi_l},
\]

(3)

In this case, only a pair of unknown parameters \((\omega, \phi)\) is estimated from the observed signal. Based on the above model, a spatio-temporal matrix model signal model is first defined as follows

\[
X(n) = S(n) + Q(n), n = 0, 1, \ldots, N - M,
\]

(4)

and \(S(n)\) and \(Q(n)\) are defined similar to \(X(n)\). In addition, \(M\) is a given positive integer, which is smaller than \(N\) and satisfies \(MR > L\). The signal matrix of interest \(S(n)\) can be rewritten by (3) as

\[
S(n) = \sum_{l=1}^{L} \beta_l(n) a_s(l\phi) a^T_s(l\omega),
\]

(6)

where

\[
\beta_l(n) = \alpha_l e^{j\omega_l n},
\]

(7)

\[
a_s(l\phi) = [e^{j\phi}, \ldots, e^{j(L-1)\phi}]^T,
\]

(8)

\[
a_s(l\omega) = [e^{j\omega_l}, \ldots, e^{j(M-1)\omega_l}]^T.
\]

(9)

Define a new vector model \(x(n)\) by stacking the columns of \(X(n)\) as

\[
x(n) = \text{vec}(X(n))
\]

\[
= s(n) + q(n) - A\{z(n)\} + q(n),
\]

(10)

and hence we have

\[
s(n) = \text{vec}(S(n))
\]

\[
q(n) = \text{vec}(Q(n))
\]

\[
A = \{a(\omega, \phi), \ldots, a(L\omega, L\phi)\},
\]

(12)

\[
a_l(\omega, \phi) = a_s(l\omega) \otimes a_s(l\phi),
\]

(13)

\[
A_l(\omega, \phi) = a_1 e^{j\omega l}, \ldots, a_L e^{jL\omega l}.
\]

(15)

Based on the data model (10), a nonlinear least squares method for DOA and pitch estimation in the case of single pitch was recently reported in [36], where the 2-D search computation as well as relatively fine candidate grids for the unknown parameters is required to obtain accurate DOA and pitch estimates, which thus results in a high computational load. In this following Section III-B, a lower-complexity ESPRIT-based method is derived based on the spatio-temporal model in the subsection, and a generalization of the proposed method to the multi-pitch case is suggested in the Section III-C.


For \(n = 0, 1, \ldots, N - M\), one first defines the following matrix with \(x[n]\),

\[
Y_{RM \times (N - M + 1)} = [x(0)x(1) \ldots x(N - M)]
\]

\[
- A [x(0) \ldots x(N - M)]
\]

\[
+ [q(0) \ldots q(N - M)]
\]

(16)

Under the condition of full column rank of matrix \(A\), one can obtain the signal subspace from the singular-value decomposition (SVD) or equivalent eigenvalue decompositions of the covariance matrix of \(Y\). The signal subspace matrix \(U_{ss}\) is then composed of the \(L\) singular vector associated with the largest singular values of \(Y\). It is well-known that the signal subspace spanned by the column vectors of \(U_{ss}\) is the same subspace as all the column vectors in matrix \(A\); therefore, there exists a unique nonsingular matrix \(T\) such that \(U_{ss} = AT\).

Using the inherent structure of \(a_{st}(\omega)\) and \(a_{st}(\phi)\), the shift invariance property of \(A\) can be found so that the classic ESPRIT method [19] is exploited to estimate the DOA and pitch parameters, respectively.

In so doing, we first define the following four selection matrices

\[
W_1 = I_{M-1} \otimes I_R,
\]

\[
W_2 = 0_{1 \times (M-1)} \otimes I_R,
\]

\[
W_3 = I_M \otimes 0_{1 \times (R-1)},
\]

\[
W_4 = I_M \otimes 0_{1 \times (R-1)} \otimes I_{H-1},
\]

(17)

where \(0\) is the vector of zeros.

The submatrices \(A_i\) of \(A\) are further defined by the above selection matrix \(W_i\) as

\[
A_i = W_i^T A, i = 1, 2, 3, 4.
\]

With the above definition and the structure of \(A\), it is easily shown that we have the following equations

\[
A_1 P(\omega) = A_2,
\]

\[
A_3 P(\phi) = A_4,
\]

(18)

(19)

where

\[
\Psi_f(\omega) = \text{diag}\{e^{j\omega}, \ldots, e^{jL\omega}\}
\]

\[
\Psi_f(\phi) = \text{diag}\{e^{j\phi}, \ldots, e^{jL\phi}\}
\]

(20)

Let \(U_1, U_2\) and \(U_3, U_4\) be the submatrices formed from \(U_{ss}\) in the same way as the \(A_i\’\s\) are formed from \(A\). Following the rationale of the ESPRIT method, the diagonal elements of \(\Psi_f(\omega)\) and \(\Psi_f(\phi)\) are the corresponding eigenvalues of \(\Phi = \...\)
\( \mathbf{T}_\psi \mathbf{T}^{-1} \) and \( \mathbf{\Phi}_s = \mathbf{T}_\psi \mathbf{T}^{-1} \), which satisfy the following equations
\[
\mathbf{U}_2 = \mathbf{U}_1 \mathbf{\Phi}_\ell \quad \text{and} \quad \mathbf{U}_4 = \mathbf{U}_3 \mathbf{\Phi}_s. \tag{21}
\]
Let \( \mu_l \) and \( \nu_l \) for \( l = 1, 2, \cdots, L \) be the corresponding eigenvalues of \( \mathbf{\Phi}_s \) and \( \mathbf{\Phi}_s \). Thus the estimates of \( \omega \) and \( \theta \) can be given as follows
\[
\hat{\omega} = \frac{2}{L(L+1)} \sum_{l=1}^{L} \mu_l, \\
\hat{\theta} = \arcsin \left( \frac{2 \nu_l \sum_{l=1}^{L} \nu_l}{L(L+1) \hat{\omega} f_d} \right). \tag{22}
\]
It can be seen from (20) that the estimation accuracy of \( \hat{\theta} \) depends on that of \( \hat{\omega} \).

C. Spatio-temporal Signal Model in the Presence of Multi-pitch

To solve the problem of joint estimation of the DOA and pitch in the multi-pitch case, a multi-channel and multi-pitch harmonic multiple signal classification (MC-HMUSIC) method was presented in [34], where a two dimensional search is required to obtain the estimates of the pitch and DOA. To further reduce the computational complexity of the 2-D search in the MC-HMUSIC method, a two-stage DOA and pitch estimation method based on a subspace approach, i.e., MCOW-HMUSIC, was recently proposed in [35]; however, the two 1-D searches and an iterative procedure for the refined estimation are still required for the MCOW-HMUSIC method. Furthermore, the coupling of the pitch and DOA parameters is still required and poses a significant problem. In this subsection, a generalization of the previously proposed ESPRIT method for joint DOA and pitch estimation of the single pitch case to the multi-pitch case is derived.

Following a procedure similar to the single pitch case, with model in (1), for a given integer \( M \) and satisfying \(RM > \sum_{p=1}^{P} \mathbf{I}_p \), one can construct the following data vector
\[
\mathbf{X}(n) = \mathbf{S}(n) + \mathbf{Q}(n), \quad n = 0, 1, \cdots, N - M, \tag{23}
\]
where
\[
\mathbf{X}(n) = \left[ \begin{array}{c}
\mathbf{x}_0(n) \\
\vdots \\
\mathbf{x}_{M-1}(n)
\end{array} \right], \tag{24}
\]
The part of the signal of interest \( \mathbf{S}(n) \) in \( \mathbf{X}(n) \) can be rewritten as
\[
\mathbf{S}(n) = \sum_{p=1}^{P} \sum_{l=1}^{L_p} \beta_{p,l}(n) \mathbf{a}_p(l \phi_p) \mathbf{a}_s^T(l \omega_p), \tag{25}
\]
where
\[
\beta_{p,l}(n) = \alpha_{p,l} e^{j\omega_{\omega_p} n}, \\
\mathbf{a}_s(\phi_p) = \left[ e^{j\phi_p}, \cdots, e^{j(R-1)\phi_p} \right]^T, \\
\mathbf{a}_s(\omega_p) = \left[ e^{j\omega_p}, \cdots, e^{j(M-1)\omega_p} \right]^T. \tag{26}
\]
Stacking the columns of \( \mathbf{X}(n) \) into a new vector model \( \mathbf{x}(n) \) gives
\[
\mathbf{x}(n) = \mathbf{v}(\mathbf{X}(n)) = \mathbf{s}(n) + \mathbf{q}(n) = \mathbf{A} \mathbf{x}(n) + \mathbf{q}(n), \tag{27}
\]
with
\[
\mathbf{A} = [\mathbf{A}(1)(\omega_1, \phi_1), \cdots, \mathbf{A}(P)(\omega_P, \phi_P)], \\
\mathbf{A}(p) = [\mathbf{a}_s(l \omega_p, \phi_p), \cdots, \mathbf{a}_s(L_p \omega_p, L_p \phi_p)], \quad p = 1, \cdots, P, \tag{28}
\]
\[
\mathbf{a}_s(l \omega_p, l \phi_p) = \mathbf{a}_s(l \omega_p) \otimes \mathbf{a}_s(l \phi_p), \quad p = 1, \cdots, P, \tag{29}
\]
\[
\mathbf{z}(n) = [\alpha_{1,1} e^{j\omega_{\omega_p}}, \cdots, \alpha_{1,L_p} e^{jL_p \omega_{\omega_p}}, \cdots, \alpha_{P,1} e^{j\omega_{\omega_p}}, \cdots, \alpha_{P,L_p} e^{jL_p \omega_{\omega_p}}]^T. \tag{30}
\]
Thus, for \( n = 0, 1, \cdots, N - M \), we first define the following vector with \( \mathbf{x}(n) \),
\[
\mathbf{Y}_{RM \times (N-M+1)} = [\mathbf{x}(0) \mathbf{x}(1) \cdots \mathbf{x}(N-M)],
\]
\[
- \mathbf{A} \mathbf{z}(0) \cdots \mathbf{z}(N-M) + [\mathbf{q}(0) \cdots \mathbf{q}(N-M)]. \tag{31}
\]

Thus, the above spatio-temporal signal model for the multi-pitch case can be reduced to the case of a single pitch in the previous section when \( P = 1 \) in (23).

D. Joint Estimation of DOA and Pitch in the Multi-pitch Case Using ESPRIT: Multichannel Multi-Pitch ESPRIT (MMP-ESPRIT)

To use the ESPRIT method for joint estimation of DOA and pitch in the multi-pitch case, the singular value decomposition (SVD) of the data matrix \( \mathbf{Y} \) is expressed by
\[
\mathbf{Y} = \mathbf{U}_s \mathbf{\Sigma} \mathbf{V}_s^H = \left[ \mathbf{U}_s \mathbf{U}_n \right] \left[ \begin{array}{cc}
\mathbf{\Sigma}_s & \mathbf{0} \\
\mathbf{0} & \mathbf{\Sigma}_s
\end{array} \right] \mathbf{V}_s^H, \tag{32}
\]
where the column vectors of \( \mathbf{U}_s \) and \( \mathbf{U}_n \) span the signal subspace and noise subspace, respectively. The diagonal elements of \( \mathbf{\Sigma}_s \) are composed of the \( P \)-largest singular values of \( \mathbf{Y} \), and one also has that \( \mathbf{U}_s \mathbf{U}_s^H = \mathbf{A} \mathbf{T} \) for a nonsingular position matrix \( \mathbf{T} \).

Using the shift invariance structure of \( \mathbf{a}_s(l \omega_p, l \phi_p) \) in \( \mathbf{A} \), we can obtain the following two equations,
\[
\mathbf{A}_1 \mathbf{\Psi}_\ell = \mathbf{A}_2, \tag{33}
\]
\[
\mathbf{A}_3 \mathbf{\Psi}_s = \mathbf{A}_4, \tag{34}
\]
where \( \mathbf{A}_i = \mathbf{W}_i^T \mathbf{A}, \quad i = 1, 2, 3, 4 \) and \( \mathbf{\Psi}_\ell \) and \( \mathbf{\Psi}_s \) are defined as
\[
\mathbf{\Psi}_\ell = \text{diag}(\mathbf{\Psi}_{\ell_1}, \cdots, \mathbf{\Psi}_{\ell_r}), \tag{35}
\]
\[
\mathbf{\Psi}_{\ell_r} = \text{diag}(\mathbf{e}^{j\phi_{\ell_r}}, \cdots, \mathbf{e}^{j(L_r-1)\phi_{\ell_r}}), \tag{36}
\]
\[
\mathbf{\Psi}_s = \text{diag}(\mathbf{\Psi}_{s_1}, \cdots, \mathbf{\Psi}_{s_p}), \tag{37}
\]
\[
\mathbf{\Psi}_{s_p} = \text{diag}(\mathbf{e}^{j\omega_{\omega_p}}, \cdots, \mathbf{e}^{j(M-1)\omega_{\omega_p}}). \tag{38}
\]
Accordingly, one can also obtain the following equations by the subspace property,
\[
\mathbf{U}_2 = \mathbf{U}_1 \mathbf{\Phi}_\ell \quad \text{and} \quad \mathbf{U}_4 = \mathbf{U}_3 \mathbf{\Phi}_s, \tag{39}
\]
where
\[
\mathbf{U}_1 = \mathbf{W}_1 \mathbf{U}_s \quad \text{and} \quad \mathbf{U}_2 = \mathbf{W}_2 \mathbf{U}_s. \tag{40}
\]
Thus, the diagonal elements of $\Phi_t$ and $\Phi_s$ can be estimated from the corresponding eigenvalues of $\Phi_t = T\Phi_sT^{-1}$ and $\Phi_s = T\Phi_tT^{-1}$, and then the estimates of pitch and DOA can be found from the eigenvalues $\mu_i^{(p)}$, $\nu_i^{(p)}$ ($i = 1, \cdots, L_p$, $p = 1, \cdots, P$) of $\Phi_t$ and $\Phi_s$, i.e.,

\begin{align}
\hat{\omega}_p &= \frac{2}{L_p (L_p + 1)} \sum_{i=1}^{L_p} \omega_i^{(p)}, \\
\hat{\theta}_p &= \arcsin \left( \frac{2c \sum_{i=1}^{L_p} \nu_i^{(p)} \omega_i f_s d}{L_p (L_p + 1) \omega_i f_s d} \right). 
\end{align}

It is worthy to note that the estimated eigenvalues need be paired with the eigenvalues of $\Phi_t$ and $\Phi_s$ so that we can determine the pairing parameters in the case of multi-pitch, however, some available methods reported in [37], [38] can be easily used to finish this pairing.

### E. Spatio-Temporal Signal Model for Single-Pitch in the Presence of Multi-path Propagation

In this subsection, the joint estimation of DOA and pitch in the presence of multi-path propagation will be considered for the single pitch case. Based on the spatio-temporally smoothed data matrix model introduced in [40], it is shown that the ESPRIT method can still be used to estimate the DOA and pitch in the multi-path propagation case.

In the signal model (2), when the multiple sources impinge on the array with different DOAs, for certain frequency combinations, the steering vector $\mathbf{u}(\omega_p, \phi_p)$ as a function of parameter pair $(\omega_p, \phi_p)$ may be the same for different $p$, especially when the multiple sources have the same pitch but distinct DOAs (i.e., for the multi-path propagation case), which causes matrix $\mathbf{A}$ to be rank-deficient in these cases. Therefore, we need to use the spatio-temporal smoothing technique [40] to restore the rank of $\mathbf{A}$. The signal model (2) for single pitch case in the presence of $P$ multi-path propagations (i.e., $\omega = \omega_1 = \cdots = \omega_P$) can be expressed as

\begin{align}
\mathbf{x}(n) &= \mathbf{s}(n) + \mathbf{q}(n), \quad n = 0, \cdots, N - 1, \\
\mathbf{s}(n) &= \sum_{p=1}^{P} \sum_{l=1}^{L_p} a_{p,l} e^{j(\omega_l n + \phi_p r_l)}.
\end{align}

The array output $\mathbf{x}(n)$ is defined as

\begin{align}
\mathbf{x}(n) &= \mathbf{x}_0(n) \mathbf{x}_1(n) \cdots \mathbf{x}_{R-1}(n)^T - \mathbf{A}\Phi_t(n) + q(n),
\end{align}

where

\begin{align}
\mathbf{A} &= [A^{(1)}, \cdots, A^{(P)}], \\
A^{(p)} &= [a(\theta_p, \omega), \cdots, a(\theta_p, L_\omega)], \\
\Phi_t(n) &= [\Phi_1(n), \cdots, \Phi_P(n)], \\
\Phi_s(n) &= \text{diag}(e^{j\omega n}, \cdots, e^{jL_\omega n}), \\
b &= [\beta_1, \cdots, \beta_L, \cdots, \beta_{P_1}, \cdots, \beta_{P_1}],
\end{align}

with the array steering vector $a(\theta_p, \omega)$ given by

\begin{align}
a(\theta_p, \omega) &= [e^{j\omega f_s \sin(\theta_p)}, \cdots, e^{j\omega f_s \sin(\theta_p) + (1 - 1) \sin(\theta_p) T}].
\end{align}

For $N$ given samples of the array output, we form the following data matrix

\begin{align}
\mathbf{X} &= [\mathbf{x}(0), \cdots, \mathbf{x}(N - 1)].
\end{align}

Using the structure of harmonic components, the above data matrix is expressed as

\begin{align}
\mathbf{X} = \mathbf{A} \mathbf{b} + \mathbf{Q}, \\
\mathbf{Q} &\in \mathbb{C}^{R \times N} \text{ is a matrix containing } N \text{ samples of the noise vector of the array output.}
\end{align}

For a given integer $M < N$, the $M$-times temporally smoothed data matrix is firstly constructed by stacking $M$ times temporally shifted versions of the original data matrix as follows [34]

\begin{align}
\mathbf{X}_M &= \\
&= \begin{bmatrix}
\mathbf{x}(0) & \mathbf{x}(1) & \cdots & \mathbf{x}(N - M) \\
\mathbf{x}(1) & \mathbf{x}(2) & \cdots & \mathbf{x}(N - M + 1) \\
\vdots & \vdots & \ddots & \vdots \\
\mathbf{x}(M - 1) & \mathbf{x}(M) & \cdots & \mathbf{x}(N - 1)
\end{bmatrix}
\end{align}

With the equation (47), $\mathbf{X}_M$ can also be written as

\begin{align}
\mathbf{X}_M &= \\
&= \mathbf{A} \mathbf{b} + \mathbf{Q}_M,
\end{align}

where $\mathbf{X}_M \in \mathbb{C}^{RM \times N-M+1}$ is the temporally smoothed data matrix and the part of the noise $\mathbf{Q}_M$ is constructed from $\mathbf{Q}$ in a similar way as $\mathbf{X}_M$ in Eq. (51).

On the other hand, $\mathbf{X}_M$ can be factorized as

\begin{align}
\mathbf{X}_M &= \\
&= \mathbf{A} \mathbf{B}_M + \mathbf{Q}_M,
\end{align}

where $\mathbf{A}_M = [\mathbf{A}^{T} \mathbf{(A}\Phi_t)^T, \cdots, (\mathbf{A}\Phi_t^{M-1})^T]^T$ and $\mathbf{B}_M = [\mathbf{b}, \mathbf{b}, \cdots, \mathbf{b}]$.

Similarly, the spatially smoothed data matrix is further constructed from the above data matrix $\mathbf{X}_M$. If we assume the number of sensors for each subarray to be $R_0$, then we can separate the array into $R - R_0 + 1$ subarrays, and the spatio-temporal smoothed data matrix $\mathbf{X}_{M,R_0} \in \mathbb{C}^{RM \times R - R_0 + 1} (N - M + 1)$ is given by

\begin{align}
\mathbf{X}_{M,R_0} &= [\mathbf{X}_{M,1}, \cdots, \mathbf{X}_{M,(R - R_0 + 1)}],
\end{align}

where

\begin{align}
\mathbf{X}_{M,r} &= \mathbf{J}_r \mathbf{X}_M = \mathbf{J}_r \mathbf{A}_M \mathbf{B}_M + \mathbf{Q}_M, r.
\end{align}

and $\mathbf{J}_r \in \mathbb{R}^{MR_0 \times MR}$ is the selection matrix corresponding to the $r$th subarray for data matrix $\mathbf{X}_M$.

It is easily shown with the shift invariance structure in $\mathbf{A}$ that the following equation holds

\begin{align}
\mathbf{A}^{(R)}_M = \mathbf{J}_r \mathbf{A}_M = \mathbf{J}_r \mathbf{A}_M \Phi_s^{r-1},
\end{align}
where

\[
\Phi_s = \text{diag}\{e^{j\phi_1}, \ldots, e^{j\phi_{LP}}, e^{j\phi_{p}}, \ldots, e^{j\phi_{LP}}\}. \tag{57}
\]

Thus, the matrix \(X_{M,R} \) can be written as

\[
X_{M,R} = J_1 A_M [B_M, \Phi_s B_M, \ldots, \Phi_s^{R-1} B_M] + Q_{M,R}, \tag{58}
\]

where \(J_1 = I_M \otimes [R_0, 0] \).


Based on the constructed spatio-temporal smoothed data matrix \(X_{M,R}, \) the noise subspace and signal subspace can be obtained from the SVD of \(X_{M,R}, \) which is defined as

\[
X_{M,R} = [U_s, U_n] \Sigma \Psi^H, \tag{59}
\]

where the columns of \(U_s \) and \(U_n \) are the singular vectors associated with the \(P_0(= LP) \) largest values and the remaining \(M R_0 - P_0 \) least singular values, respectively. Using the noise subspace of \(X_{M,R}, \) a MUSIC-based method for joint multi-pitch and DOA estimation was presented in [34], where a 2-D grid search is required to obtain the estimates of parameter pairs \((\omega, \theta_p), p = 1, 2, \ldots, P. \) In this section, using the spatio-temporal shift invariance structure of \(A_{M,R}, \) the ESPRIT-based method is exploited to give the estimates of \(\{\omega, \theta_p\}. \) One can obtain the following two equations with the signal subspace matrix \(U_s, \)

\[
U_2 = U_1 \Phi_s \quad \text{and} \quad U_4 = U_3 \Phi_s, \tag{60}
\]

where

\[
U_1 = W_1 U_s \quad \text{and} \quad U_2 = W_2 U_s, \quad U_3 = W_3 U_s \quad \text{and} \quad U_4 = W_4 U_s, \tag{61}
\]

where the selection matrices \(W_i, i = 1, 2, 3, 4, \) are defined as follows

\[
W_1 = \begin{bmatrix} I_{M-1} & 0_{1 \times (M-1)} \end{bmatrix} \otimes I_{R_0}, \quad W_2 = \begin{bmatrix} 0_{1 \times (M-1)} & I_{M-1} \end{bmatrix} \otimes I_{R_0}, \quad W_3 = I_M \otimes [I_{R_0-1} 0_{1 \times (R_0-1)}], \quad W_4 = I_M \otimes [0_{1 \times (R_0-1)} I_{R_0-1}].
\]

Thus, the estimates of \(\omega \) and \(\theta_p \) can be obtained from the eigenvalues \(\mu_i \) and \(\mu_i^{(p)} \) of \(\Phi_s \) and \(\Phi_s^{(p)} \) as

\[
\hat{\omega} = \frac{2}{L(L+1)} \sum_{i=1}^{L} \omega_i, \quad \hat{\theta}_p = \text{arcsin} \left( \frac{2 \sum_{i=1}^{L} \mu_i^{(p)} \omega_i f_i d}{L(L+1) \omega^2 f_i d^2} \right), \quad p = 1, 2, \ldots, P. \tag{62}
\]

### IV. ALGORITHM SUMMARIZATION AND COMPUTATIONAL COMPLEXITY ANALYSIS

In this section we analyze the computational complexity of the proposed algorithms. To that end, Table I provides a detailed description of the three algorithms.

<table>
<thead>
<tr>
<th>Algorithm</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>MSP-ESPRIT</td>
<td>1. Align the observed matrix ( Y ) in Eq.(16)</td>
</tr>
<tr>
<td></td>
<td>2. Compute the SVD of ( Y )</td>
</tr>
<tr>
<td></td>
<td>3. Compute the eigenvalues of the two matrices ( \Phi_1 ) and ( \Phi_s ) in Eq.(21)</td>
</tr>
<tr>
<td></td>
<td>4. Estimate the pitch and DOA using Eqs.(22)</td>
</tr>
<tr>
<td>MMP-ESPRIT</td>
<td>1. Align the observed matrix ( Y ) in Eq.(31)</td>
</tr>
<tr>
<td></td>
<td>2. Compute the SVD of ( Y ) in Eq.(32)</td>
</tr>
<tr>
<td></td>
<td>3. Compute the eigenvalues of the two matrices ( \Phi_1 ) and ( \Phi_s ) in Eq.(37)</td>
</tr>
<tr>
<td></td>
<td>4. Estimate the pitch and DOA using Eqs.(39)</td>
</tr>
<tr>
<td>MSP-MP-ESPRIT</td>
<td>1. Align the observed matrix ( X_{M,R_0} ) in Eq.(58)</td>
</tr>
<tr>
<td></td>
<td>2. Compute the SVD of matrix ( X_{M,R_0} ) in Eq.(59)</td>
</tr>
<tr>
<td></td>
<td>3. Compute the eigenvalues of the two matrices ( \Phi_1 ) and ( \Phi_s ) in Eq.(60)</td>
</tr>
<tr>
<td></td>
<td>4. Estimate the pitch and DOA using Eqs.(62)</td>
</tr>
</tbody>
</table>

### V. ASYMPTOTIC PERFORMANCE ANALYSIS OF PITCH AND DOA ESTIMATES

In this section, the asymptotic performance of the proposed ESPRIT-based method for multi-pitch and DOA estimation is analyzed. With the similar results derived in [42], it is easily shown that the estimations of pitch and DOA parameters are unbiased. The derivation of this result is based on the same analysis procedure for DOA estimation using the ESPRIT method in [41], while it can be directly generalized to the case of multi-path propagation scenarios, therefore, The derivation of this result is based on the same analysis procedure for DOA estimation using the ESPRIT method in [41].

We provide the variances of the estimated pitch and DOA parameters for the multi-pitch case in the absence of multi-path propagation, which are denoted by \( E\{\delta \omega_k^2\} \) and \( E\{\delta \theta_p^2\}. \)
TABLE II
SIMULATION SETTING OF THE MSP-ESPRIT, MC-HMUSIC
AND THE NLS-BASED METHOD

<table>
<thead>
<tr>
<th>p</th>
<th>L</th>
<th>Pitch</th>
<th>DOA</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>5</td>
<td>243</td>
<td>15</td>
</tr>
</tbody>
</table>

The variances are based on the covariance of the eigenvalues of \( \mu_1^{(p)}, \nu_1^{(p)} \), \( p = 1, \cdots, P \).

The detailed derivations are found in the Appendix, and yield

\[
E\{ \delta \omega_2^p \} = \frac{1}{L_P} E\left\{ \sum_{l=1}^{L_P} \text{real}\{ j \mu_1^{(p)} \delta \mu_1^{(p)} \} \right\}^2, \tag{63}
\]

and

\[
E\{ \delta \theta_2^p \}^2 = \frac{1}{(L_P \omega_2 f_s \frac{d}{c} \cos \theta_2)^2} E\{ \sum_{l=1}^{L_P} \text{real}\{ -j \nu_1^{(p)} \delta \nu_1^{(p)} + j \mu_1^{(p)} \frac{d}{c} \sin \theta_2 \delta \mu_1^{(p)} \} \}^2. \tag{64}
\]

VI. SIMULATION RESULTS

To evaluate the performance of the proposed method for joint estimation of the pitch and DOA, a set of Monte Carlo simulations is presented using synthetic data followed by experiments on real-life data as presented in the final subsection.

A. Synthetic Data Results

The first evaluation was conducted on a set of simulations composed of synthetic data. In all the simulations, it is assumed that the array is a uniform linear array with inter-element spacing \( d = c/f_s \), where the speed of sound is set as \( c = 343.2 \) m/s and the sampling frequency is \( f_s = 8 \) kHz. All the amplitudes of the sinusoids are assumed to be units, i.e., \( \beta_{p,l} = 1 \) for each \( (p, l) \). The parameter \( M = N/2 \) was used in all the experiments. The estimation accuracy was evaluated using the root mean square error (RMSE): \( \text{RMSE}_{\omega} = \sqrt{\frac{1}{Q} \sum_{q=1}^{Q} (\hat{\omega}_q - \omega)^2} / Q \) and \( \text{RMSE}_{\theta} = \sqrt{\frac{1}{Q} \sum_{q=1}^{Q} (\hat{\theta}_q - \theta)^2} / Q \), where \( \omega \) and \( \theta \) are the true pitch and DOA, and \( \hat{\omega} \) and \( \hat{\theta} \) are their estimates, respectively. Furthermore, \( Q = 200 \) is the number of independent trials. In all the simulations, the signal-to-noise ratio (SNR) was defined as

\[
\text{SNR} = 10 \log_{10} \frac{\sum_{p=1}^{P} \sum_{l=1}^{L_p} |\beta_{p,l}|^2}{\sigma^2}
\]

In the first experiment, we first provide an example of single-pitch estimation. The number of microphones was set to \( R = 4 \) and the number of samples was \( N = 80 \). The other experimental parameter setting is listed in Table II. For the purpose of comparison, the results obtained using the NLS-based method reported in [36] as well the MC-HMUSIC method, and the corresponding CRLB were also included in all the simulation results. Since it was shown in [36] that the performance of the NLS-based method is optimal out of all the existing methods such as LCMV [32] and MC-ML method [28], we only compare the performance of the proposed method with those of both the NLS-based method and the MC-HMUSIC.

The RMSEs of the proposed estimators (MSP-ESPRIT) versus different SNRs were compared to those of both the NLS-based method and the MC-HMUSIC method. Fig. 1 shows that the performance of the proposed method is very similar to the latter two methods and also follows the corresponding CRLB in all SNRs cases, and that the corresponding theoretical values of the proposed estimation fit the simulation results. We also observe from Fig. 1 that the performance in terms of the RMSE, seem to be below the CRB in some cases (SNR = 25 dB). However, this is due to the limited number of Monte Carlo trials, making the measured RMSE fluctuate around the CRB. In addition, we fix the number of samples as \( N = 100 \) and check the average elapsed time of three different methods for single run versus different number of sensors. It is seen from Fig. 2 that the simulation result also confirms the significant reduction in computational complexity obtained with the proposed approach.

In the second experiment, we evaluated estimations in the multi-pitch case. The harmonic signal consisted of \( P = 2 \) pitches, each with \( L_1 = L_2 = 2 \) tones. The other experimental
parameter setting is listed in Table III. The number of sensors was \( R = 8 \) and the number of samples was \( N = 120 \), respectively. In the experiment, the performance of the recently presented MC-HMUSIC method [34] was compared to that of the proposed method. Figs. 3–4 shows the RMSEs for pitch and DOA estimates of the two methods in the different SNR cases. Figs. 3–4 indicates that the performance of the proposed method is similar to that of the MC-HMUSIC estimator and the RMSEs of the two methods are also close to the corresponding CRLB for all SNRs. Furthermore, the simulated results are consistent with those of the theoretical variances in (63) and (64).

For the purpose of further comparison, the resolvability performance in frequency as well as the angle of the two methods was compared. The SNR was set to 25 dB and the number of samples was \( N = 120 \). The other parameters in this experiment were the same as those in the second experiment. Firstly, both the parameters of source 1 and the DOA of source 2 were fixed, while the pitch of source 2 varied with different \( \Delta \omega = |\omega_1 - \omega_2| \). Fig. 5 shows that the resolvability performance of
the two methods is similar, but there is about a 2 dB gap between their RMSEs and the corresponding CRLB.

Similar to the above experiment on resolvability performance in frequency, both parameters of source 1 and the pitch of source 2 were now fixed in this experiment, while the DOA of source 2 varied with different $\Delta \theta = |\theta_1 - \theta_2|$. It is seen from the Fig. 6 that the resolvability performance in spatial angle of the two methods is also similar, but with a 2 dB gap between their RMSEs and the corresponding CRLB. In addition, it is worthy to note that both the MC-HMUSIC estimator and the proposed estimators can still perform well when the DOAs of the two sources are identical, i.e., $\Delta \theta = 0$, due to the fact that the two sources have different pitches and the estimations of DOA are dependent on the pitch parameters. However, the pitch estimates of the two sources are directly obtained from the eigenvalues of constructed matrices; therefore, the two sources cannot be resolved in the case of the same pitch (i.e., $\Delta \omega = 0$) without using the spatio-temporal smoothing technique to restore the rank of $\mathbf{A}$.

In the last simulation experiment, we show an example of joint DOA and single-pitch estimation in the presence of multi-path propagation. The harmonic signal consisted of $L$ sinusoids with a pitch of $f_0 = 343$ Hz and assumed that there were 3 multi-path signals with DOAs $= [-60^\circ, 10^\circ, 45^\circ]$ received at the array, which the parameter setting is listed in Table IV. The number of microphones was set $R = 16$, SNR was set as $\text{SNR} = 30$ dB, and the number of samples was $N = 120$. The smoothed parameters in the spatio-temporal domain were taken as $M = 80$ and $R_0 = 8$. The results for 50 independent trials are shown in Fig. 7, which shows that the proposed method can still give accurate DOA and pitch estimation in the multi-path propagation case, and that the result for $L = 5$ is better than the result for $L = 2$ in same conditions.

### B. Real-life Data Results

In this subsection, an experiment was conducted to also evaluate the performance of the proposed method on real-life signals. The real-life data were also used to measure the performance of the NLS-based method [36] for the single pitch case. The experiment took place in a meeting room and the experimental environment was depicted in [36]. In the experiment, the sampling frequency was $f_s = 44.1$ kHz, the speed of sound was assumed to be $c = 343.2$ m/s, and the room reverberation time at 1 kHz was $T_{60} \approx 0.53$ s, the receiving array was a uniform and linear array with an inter-element spacing of $d = 4$ cm and the number of sensors was $R = 8$. We assumed that there was no attenuation across the sensors such that $\alpha_{p,i} = 1$, and the desired signal was assumed to consist of $L = 8$ harmonics. In addition, $N = 10^6$ was used for each parameter estimation.
The estimates in the experiment were set up as in the previous simulations with synthetic data.

In the part of the real-life experiment, the single pitch case was tested using a played-back trumpet note “S1” in [36]. The anechoic trumpet signal was generated by concatenating anechoic trumpet signal excerpts (see http://theremin.music.uiowa.edu/MIS.html) [43]. The played back trumpet signal was recorded using the ULA to obtain a multichannel trumpet signal with slight reverberation, and the DOA of the single trumpet signal was about $\theta_1 = -13^\circ$. From the recorded data, we estimated the pitch and the DOA of the trumpet signal via the proposed method and the NLS-based method. The results obtained from this experiment are shown in Fig. 9. It can be seen that the two applied estimators for the pitch estimation yielded identical estimates in the recorded time interval, but that the estimation accuracy of the pitch using the NLS-based method was better than the proposed method, when the results of the pitch parameter were only compared to the corresponding spectrogram shown in Fig. 8. However, the estimated DOAs using the two methods were relatively close to the true DOA.

**VII. CONCLUSION**

In this paper, the ESPRIT algorithm was used for joint pitch and DOA estimation of multiple channel multi-pitch with a ULA. Based on a constructed spatio-temporal signal model, the invariance property in both the time-domain and the spatial domain was exploited to estimate the pitch and DOA parameters using the ESPRIT method. We also demonstrated that the spatio-temporal smoothing technique can be used to solve the single-pitch multi-path propagation case. Compared to the available state-of-the-art algorithms, the proposed method based on ESPRIT without 2-D searching is computationally more efficient. The asymptotic performance of the estimated pitch and DOA for multi-pitch case was also analyzed. Finally, numerous simulations using synthetic data were conducted to compare the performance of the proposed method with both the NLS-based method and the MC-HMUSIC method for the single pitch case and the MC-HUMUSIC method for the multi-pitch case. The results show that the performance of the proposed estimator is not only comparable to the other methods for both pitch and DOA estimation in terms of the root mean square error but is also close to the corresponding CRLB. Furthermore, the experimental test on real-life data demonstrates that the
proposed method can be applied to real-life signals and yields a similar estimation performance to the NLS-based method. In future work, the problem of joint DOA and pitch estimation for multi-pitch and multi-path scenarios will be considered.

APPENDIX

Following the steps in [41], the derivation of the variances of the estimated pitch and DOA is given as follows:

(1) Variance of pitch estimation: For the pitch estimation, the parameter $\hat{\omega}_p$ is related to the eigenvalue $\mu^{(p)}_1$ of $\Phi$ by

$$\mu^{(p)}_1 = e^{j\omega_p l}, l = 1, \ldots, L_p. \quad (65)$$

Using the first-order Taylor series expansion, yields

$$\delta \mu^{(p)}_1 = j \mu^{(p)}_1 l \delta \omega_p, l = 1, \ldots, L_p. \quad (66)$$

Thus, we have

$$\delta \omega_p = \frac{1}{\mu^{(p)}_1 l} \delta \mu^{(p)}_1. \quad (67)$$

For $l = 1, \ldots, L_p$, we take the average of all the $\delta \omega_p$,

$$\delta \omega_p = \frac{1}{L_p} \sum_{l=1}^{L_p} \frac{1}{\mu^{(p)}_1 l} \delta \mu^{(p)}_1. \quad (68)$$

Since the $\delta \omega_p$ is a real quantity, we get

$$\delta \mu^{(p)}_1 = \frac{1}{2L_p} \sum_{l=1}^{L_p} \text{real}\left\{ \frac{1}{\mu^{(p)}_1 l} \delta \mu^{(p)}_1 + \frac{1}{j \mu^{(p)}_1 l} \delta \mu^{(p)}_1 \right\}$$

$$= \frac{1}{L_p} \sum_{l=1}^{L_p} \text{real}\left\{ -j \mu^{(p)}_1 l \delta \mu^{(p)}_1 \right\} \quad (69)$$

The variance of the estimated pitch parameter is

$$E\{\{\delta \omega_p\}^2\} = \frac{1}{L_p^2} E\{\sum_{l=1}^{L_p} \text{real}\left\{ j \mu^{(p)}_1 l \delta \mu^{(p)}_1 \right\} \}^2 \quad (70)$$

For simplicity, here we consider two harmonic components with $L_p = 2$ for $E\{\{\omega_r\}^2\}$ and $E\{\{\theta_r\}^2\}$, respectively.

When $L_p = 2$, we have

$$\delta \omega_p = \frac{1}{2} \left( \frac{1}{j \mu^{(p)}_1} \delta \mu^{(p)}_1 + \frac{1}{2j \mu^{(p)}_2} \delta \mu^{(p)}_2 \right). \quad (71)$$

Since the $\delta \omega_p$ is a real quantity, we get

$$\delta \omega_p = \frac{1}{2} (\delta \omega_p + \delta \omega_p^*) \quad (72)$$

After complicated algebra, the variance of the estimated pitch parameter is

$$E\{\{\delta \omega_p\}^2\} = \frac{1}{2} \left( \delta \omega_p + \delta \omega_p^* \right) \left( \delta \omega_p + \delta \omega_p^* \right)$$

$$= \frac{1}{8} \left| \mu^{(p)}_1 \right|^2 E\{\{\delta \mu^{(p)}_1\}^2\}$$

$$+ \frac{1}{32} \left| \mu^{(p)}_2 \right|^2 E\{\{\delta \mu^{(p)}_2\}^2\}$$

$$- \frac{1}{8} \text{real} \left( \mu^{(p)}_1 \delta \mu^{(p)}_1 \right) E\{\{\delta \mu^{(p)}_1\}^2\}$$

$$\quad - \frac{1}{8} \text{real} \left( \mu^{(p)}_2 \delta \mu^{(p)}_2 \right) E\{\{\delta \mu^{(p)}_2\}^2\}$$

$$- \frac{1}{8} \left( \mu^{(p)}_2 \delta \mu^{(p)}_2 \right) E\{\{\delta \mu^{(p)}_1\}^2\}$$

$$- \frac{1}{8} \left( \mu^{(p)}_1 \delta \mu^{(p)}_1 \right) E\{\{\delta \mu^{(p)}_2\}^2\} \quad (73)$$

(2) DOA: For DOA estimation, the parameter $\theta_p$ is dependent on $\omega_p$, which is related to $\nu^{(p)}_l$ by

$$\nu^{(p)}_l = e^{j\omega_p f_s \pi \sin \theta_p}, l = 1, \ldots, L_p. \quad (74)$$

Using the first-order Taylor series expansion and (40), after some algebra, yields

$$\delta \theta_p = \frac{\delta \nu^{(p)}_l - j \nu^{(p)}_l l f_s \pi \sin \theta_p \delta \omega_p}{j \nu^{(p)}_l l \omega_p f_s \pi \cos \theta_p} \quad (75)$$

For $l = 1, \ldots, L_p$, we take the average of all the $\delta \theta_p$, and get

$$\delta \theta_p = \frac{1}{L_p} \sum_{l=1}^{L_p} \frac{\delta \nu^{(p)}_l - j \nu^{(p)}_l l f_s \pi \sin \theta_p \delta \omega_p}{j \nu^{(p)}_l l \omega_p f_s \pi \cos \theta_p}$$

$$= \frac{1}{L_p \omega_p f_s \pi \cos \theta_p} \sum_{l=1}^{L_p} \left\{ -j \nu^{(p)}_l \delta \nu^{(p)}_l \right\}$$

$$+ j \nu^{(p)}_l f_s \pi \sin \theta_p \delta \omega_p \quad (76)$$

Similarly, we take

$$\delta \theta_p = \frac{1}{L_p \omega_p f_s \pi \cos \theta_p} \sum_{l=1}^{L_p} \left\{ -j \nu^{(p)}_l \delta \nu^{(p)}_l \right\}$$

$$+ j \nu^{(p)}_l f_s \pi \sin \theta_p \delta \omega_p \quad (77)$$

and, hence, the variance of the estimated DOA parameter is computed as

$$E\{\{\delta \theta_p\}^2\} = \frac{1}{L_p \omega_p f_s \pi \cos \theta_p} \sum_{l=1}^{L_p} \left\{ -j \nu^{(p)}_l \delta \nu^{(p)}_l \right\}$$

$$+ j \nu^{(p)}_l f_s \pi \sin \theta_p \delta \omega_p \quad (78)$$

Similarly, for $L_p = 2$, we have

$$\delta \theta_p = \frac{1}{2L_p \omega_p f_s \pi \sin \theta_p} \sum_{l=1}^{2} \left\{ -j \nu^{(p)}_l \delta \nu^{(p)}_l \right\}$$

$$+ j \nu^{(p)}_l f_s \pi \sin \theta_p \delta \omega_p \quad (79)$$

where $A = \omega_p f_s \pi \sin \theta_p$ and $H = f_s \pi \sin \theta_p$.

Let $\delta \theta_p = 2(\delta \omega_p + \delta \omega_p^*)$, and then we have the equation shown at the bottom of the next page.
Using a similar derivation similar to the one in [41], the quantities \( E\{\delta \mu_1^{(p)}\} \), \( E\{\delta \mu_2^{(p)}\} \), \( E\{\delta \nu_1^{(p)}\} \), \( E\{\delta \nu_2^{(p)}\} \), \( E\{\delta \mu_1^{(*)}\} \), \( E\{\delta \mu_2^{(*)}\} \) are the covariances of eigenvalues, which are derived one by one.

\[
\begin{align*}
\delta \mu_1^{(p)} &= u_1^{(p)} q_1^{(p)} (W_2 - \mu_1^{(p)} W_1) \delta U_s q_1^{(p)}, \\
\delta \mu_2^{(p)} &= u_2^{(p)} q_2^{(p)} (W_2 - \mu_2^{(p)} W_1) \delta U_s q_2^{(p)}, \\
\delta \nu_1^{(p)} &= u_1^{(p)} q_1^{(p)} (W_2 - \mu_1^{(p)} W_1) \delta U_s h_1^{(p)}, \\
\delta \nu_2^{(p)} &= u_2^{(p)} q_2^{(p)} (W_2 - \mu_2^{(p)} W_1) \delta U_s h_2^{(p)}.
\end{align*}
\]  

(81)

where \( u_1^{(p)}, q_1^{(p)} \) are the corresponding left and right eigenvectors of \( \Phi \), and \( \delta U_s = U_s - U_s \), which is given as follows (see [42] for details),

\[
\delta U_s = U_n U_n^H \delta Y V_s \Sigma_s^{-1},
\]

(82)

\[
\delta Y = Y - \bar{Y}.
\]

(83)

Similarly, \( v_i^{(p)}, h_i^{(p)} \) are the corresponding left and right eigenvectors of the matrix \( \Phi_s \), and the error of \( \delta \nu_i \) is then given by

\[
\delta \nu_i^{(p)} = v_i^{(p)} u_i^H (W_2 - v_i^{(p)} W_1) \delta U_s h_i^{(p)}.
\]

(84)

From (81) and (84), the variances of \( \mu_1^{(p)} \) and \( \mu_2^{(p)} \) are thus computed by the following expressions

\[
E\{\delta \mu_1^{(p)}\} = u_1^{(p)} U_1^H (W_2 - \mu_1^{(p)} W_1) E\{\delta U_s q_1^{(p)}
\]

(85)

\[
= u_1^{(p)} U_1^H (W_2 - \mu_1^{(p)} W_1) U_n U_n^H E\{\delta Y V_s \Sigma_s^{-1} q_1^{(p)} q_1^{(p)\dagger} V_s^{-1} \delta Y^H \}
\]

\[
= U_n U_n^H (W_2 - \mu_1^{(p)} W_1)^H (u_1^{(p)} U_1^H)^H
\]

\[
E\{\delta \nu_i^{(p)}\} = v_i^{(p)} U_{i_1}^H (W_2 - v_i^{(p)} W_1) U_n U_n^H E\{\delta Y V_s \Sigma_s^{-1} q_i^{(p)} q_i^{(p)\dagger} V_s^{-1} \delta Y^H \}
\]

(86)

\[
E\{\delta \mu_1^{(*)}\} = u_1^{(*)} U_1^H (W_2 - \mu_1^{(*)} W_1) \delta U_s q_1^{(p)}
\]

(87)

\[
E\{\delta \nu_i^{(*)}\} = u_i^{(*)} U_i^H (W_2 - v_i^{(*)} W_1) \delta U_s h_i^{(p)}
\]

(88)

\[
E\{\delta \mu_1^{(p)}\} = u_1^{(p)} U_1^H (W_2 - \mu_1^{(p)} W_1) \delta U_s q_1^{(p)}
\]

(89)

where \( u_i^{(*)} \) is the \( i \)-th element of the vector \( u_1^{(*)} \), \( v_i^{(*)} \) is the \( i \)-th element of the vector \( v_i^{(*)} \), and \( \delta \nu_i \) is the \( i \)-th element of the vector \( \delta \nu_i \).

In the equation (85), let \( l_1^{(p)} = V_s \Sigma_s^{-1} q_1^{(p)} \) and \( g_1^{(p)} = V_s \Sigma_s^{-1} h_1^{(p)} \), and then we get

\[
E\{\delta Y V_s \Sigma_s^{-1} q_1^{(p)} q_1^{(p)\dagger} V_s^{-1} \delta Y^H \}
\]

(90)

\[
E\{\delta Y V_s \Sigma_s^{-1} l_1^{(p)} l_1^{(p)\dagger} V_s^{-1} \delta Y^H \}
\]

(91)

\[
E\{\delta \nu_i^{(p)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(p)}\}
\]

(92)

\[
E\{\delta \nu_i^{(*)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(*)}\}
\]

(93)

\[
E\{\delta \nu_i^{(p)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(p)}\}
\]

(94)

\[
E\{\delta \nu_i^{(*)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(*)}\}
\]

(95)

\[
E\{\delta \nu_i^{(p)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(p)}\}
\]

(96)

\[
E\{\delta \nu_i^{(*)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(*)}\}
\]

(97)

\[
E\{\delta \nu_i^{(p)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(p)}\}
\]

(98)

\[
E\{\delta \nu_i^{(*)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(*)}\}
\]

(99)

\[
E\{\delta \nu_i^{(p)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(p)}\}
\]

(100)

\[
E\{\delta \nu_i^{(*)}\} = \frac{\lambda_i^{(*)} \lambda_i^{(*)}}{\lambda_i^{(*)} \lambda_i^{(*)}} E\{\delta \nu_i^{(*)}\}
\]

(101)
matrix $\delta Y$. The simplified formulas in Equations (86)-(89) can be derived similarly and are omitted here.

**REFERENCES**


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