

Incidence and predictors of atrial fibrillation episodes as detected by implantable loop recorder in patients at risk From the LOOP study

Diederichsen, Søren Zöga; Haugan, Ketil Jørgen; Brandes, Axel; Graff, Claus; Krieger, Derk; Kronborg, Christian; Holst, Anders Gaarsdal; Nielsen, Jonas Bille; Køber, Lars; Højberg, Søren; Svendsen, Jesper Hastrup

Published in:
American Heart Journal

DOI (link to publication from Publisher):
[10.1016/j.ahj.2019.09.009](https://doi.org/10.1016/j.ahj.2019.09.009)

Creative Commons License
CC BY-NC-ND 4.0

Publication date:
2020

Document Version
Accepted author manuscript, peer reviewed version

[Link to publication from Aalborg University](#)

Citation for published version (APA):
Diederichsen, S. Z., Haugan, K. J., Brandes, A., Graff, C., Krieger, D., Kronborg, C., Holst, A. G., Nielsen, J. B., Køber, L., Højberg, S., & Svendsen, J. H. (2020). Incidence and predictors of atrial fibrillation episodes as detected by implantable loop recorder in patients at risk From the LOOP study. *American Heart Journal*, 219, 117-127. <https://doi.org/10.1016/j.ahj.2019.09.009>

General rights

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
- You may not further distribute the material or use it for any profit-making activity or commercial gain
- You may freely distribute the URL identifying the publication in the public portal -

Take down policy

If you believe that this document breaches copyright please contact us at vbn@aub.aau.dk providing details, and we will remove access to the work immediately and investigate your claim.

Downloaded from vbn.aau.dk on: December 05, 2025

Incidence and predictors of atrial fibrillation episodes as detected by implantable loop recorder in patients at risk From the LOOP study

Søren Zöga Diederichsen, Ketil Jørgen Haugan, Axel Brandes, Claus Graff, Derk Krieger, Christian Kronborg, Anders Gaarsdal Holst, Jonas Bille Nielsen, Lars Køber, Søren Højberg, Jesper Hastrup Svendsen



PII: S0002-8703(19)30267-4

DOI: <https://doi.org/10.1016/j.ahj.2019.09.009>

Reference: YMHJ 5983

To appear in: *American Heart Journal*

Received date: 6 July 2019

Accepted date: 8 September 2019

Please cite this article as: S.Z. Diederichsen, K.J. Haugan, A. Brandes, et al., Incidence and predictors of atrial fibrillation episodes as detected by implantable loop recorder in patients at risk From the LOOP study, *American Heart Journal*(2019), <https://doi.org/10.1016/j.ahj.2019.09.009>

This is a PDF file of an article that has undergone enhancements after acceptance, such as the addition of a cover page and metadata, and formatting for readability, but it is not yet the definitive version of record. This version will undergo additional copyediting, typesetting and review before it is published in its final form, but we are providing this version to give early visibility of the article. Please note that, during the production process, errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

Incidence and predictors of atrial fibrillation episodes as detected by implantable loop recorder in patients at risk

From the LOOP study

Short title: Atrial fibrillation detected by loop recorder

Authors and affiliations:

Søren Zöga DIEDERICHSEN, MD^a, Ketil Jørgen HAUGAN, MD, PhD^b, Axel BRANDES, MD, DMSc^{c,d}, Claus GRAFF, PhD^e, Derk KRIEGER, MD, PhD^{f,g}, Christian KRONBORG, PhD^h, Anders Gaarsdal HOLST, MD, PhDⁱ, Jonas Bille NIELSEN, MD, PhD^{j,k}, Lars KØBER, MD, DMSc^{a,l}, Søren HØJBERG, MD, PhD^m, Jesper Hastrup SVENDSEN, MD, DMSc^{a,i,l}.

^a Department of Cardiology, Rigshospitalet, Copenhagen University Hospital
Address: Blegdamsvej 9, 2100 Copenhagen, DENMARK

^b Department of Cardiology, Sjælland University Hospital Roskilde
Address: Sygehusvej 10, 4000 Roskilde DENMARK

^c Department of Cardiology, Odense University Hospital
Address: J. B. Winsløws Vej 4, 5000 Odense, DENMARK

^d Department of Clinical Research, Faculty of Health Sciences, University of Southern Denmark
Address: Winsløwparken 19, 5000 Odense C, DENMARK

^e Department of Health Science and Technology, Aalborg University
Address: Fredrik Bajers Vej 7 D2, 9220 Aalborg, DENMARK

^f University Hospital Zurich, University of Zurich
Address: Rämistrasse 100, 8091 Zürich, SWITZERLAND

^g Stroke Unit, Mediclinic City Hospital
Address: Building 37 - 26th St, Dubai, UAE

^h Department of Business and Economics, University of Southern Denmark
Address: Campusvej 55, 5230 Odense, DENMARK

ⁱ Laboratory for Molecular Cardiology, Department of Cardiology, Rigshospitalet, Copenhagen University Hospital
Address: Blegdamsvej 9, 2100 Copenhagen, DENMARK

^j Department of Epidemiology Research, Statens Serum Institut
Address: Artillerivej 5, 2300 Copenhagen, DENMARK

^k K.G. Jebsen Center for Genetic Epidemiology, Faculty of Medicine and Health Sciences, Norwegian University of Science and Technology
Address: Håkon Jarls Gate 11, 7491 Trondheim, NORWAY

^l Department of Clinical Medicine, Faculty of Health and Medical Sciences, University of Copenhagen
Address: Blegdamsvej 3B, 2200 Copenhagen, DENMARK

^m Department of Cardiology, Bispebjerg Hospital, Copenhagen University Hospital
Address: Bispebjerg Bakke 23, 2400 Copenhagen DENMARK

Word count: 3997 (body text)
Word count, full manuscript: 6352 (highlights + body text + references + figure legends + Table 1)
Number of tables: 1 (excluding 3 supplementary tables)
Number of figures: 5 (excluding 2 supplementary figures)

Corresponding author:

Søren Zöga Diederichsen, MD
Department of Cardiology, The Heart Centre, Rigshospitalet, Copenhagen University Hospital
Blegdamsvej 9, 2100 København, DENMARK
Mail: Soeren.Zoega.Diederichsen@regionh.dk
Phone: +45 23450489 Cell: +45 2345 0489 Fax: +45 3545 2705

ABSTRACT

Background: Recent studies have suggested a high prevalence of subclinical atrial fibrillation (AF) in various patient populations, and interest in AF screening has increased. However, knowledge about episode-duration is scarce, and risk factors for short or long subclinical AF episodes have yet to be recognized.

Aims: To assess AF by long-term continuous screening, and to investigate predictors of episodes lasting ≥ 6 minutes, ≥ 5.5 hours or ≥ 24 hours, respectively.

Methods: A total of 597 patients aged ≥ 70 years and diagnosed with ≥ 1 of hypertension, diabetes, previous stroke, or heart failure, were recruited from the general population to receive implantable loop recorder with remote monitoring. Exclusion criteria included history of AF or cardiac implantable electronic device. AF episodes were adjudicated by senior cardiologists.

Results: During 40 [37;42] months of continuous monitoring, AF was detected in 209 (35%) of the patients. The cumulative incidences at 3 years were 33.8 (30.2-37.8), 16.1 (13.4-19.4), and 5.7 (4.1-7.9) % for AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours and ≥ 24 hours, respectively. Slower resting sinus rate and higher body mass index, NT-proBNP, and troponin T at baseline were independently associated with AF detection. Addition of these markers to a model of sex, age, and comorbidities improved prediction of AF episodes ≥ 24 hours (time-dependent area under the receiver operating characteristic curve 79% vs. 65%, $p=0.037$).

Conclusion: A considerable burden of previously unknown AF was detected when long-term monitoring was applied in at-risk patients. Biomarkers were associated with AF incidence and improved prediction of long AF episodes.

Keywords: Atrial fibrillation; Implantable loop recorder; Remote monitoring; Screening; Risk factors

Highlights

- 35% of at-risk patients had new-onset AF during long-term continuous screening
- Episodes lasting ≥ 5.5 and ≥ 24 hours were seen in 16 and 6 %, respectively
- A model of sex, age, and comorbidities poorly identified participants at risk of AF
- BMI, heart rate, NT-proBNP, and troponin T improved prediction of long AF episodes
- These biomarkers could be used to select patients for screening

Journal Pre-proof

INTRODUCTION

Ischemic stroke is an increasing health problem world-wide.¹ At least 20% of ischemic strokes are attributable to atrial fibrillation (AF),² and another 30% of ischemic strokes are cryptogenic, possibly related to undiagnosed AF.³ Oral anticoagulation (OAC) is well established as an effective treatment for stroke prevention in at-risk patients diagnosed with AF.⁴ However, as AF is often asymptomatic, many patients remain undiagnosed. Approximately 30% of a general population of pacemaker or cardioverter defibrillator patients, will have new-onset AF during 2-3 years following implantation.^{3,5} Although the majority of these AF episodes are short-lasting and asymptomatic, the *ASymptomatic atrial fibrillation and Stroke Evaluation in pacemaker patients and the atrial fibrillation Reduction atrial pacing Trial* (ASSERT) and other studies have found that such AF episodes are associated with risk of stroke.^{5,6} Since the ASSERT was published, screening for AF has received increasing attention.⁷ Still, the bulk of our knowledge about AF and stroke risk is derived from patients with symptomatic or rather long-lasting AF episodes, and mass screening cannot yet be recommended.^{4,7} One problem is that an appropriate screening population has yet to be recognized. A post-hoc analysis of ASSERT found that the increased risk of stroke was driven by patients with AF episodes lasting >24 hours,⁸ and a recent consensus document suggested the compromise of OAC in patients with device-detected AF episodes lasting ≥5.5 hours.⁹ Similarly, recent data from the Veterans Health Administration showed increasing rates of stroke with increasing episode duration.¹⁰ Predictors for such AF episodes remain unknown.

The ongoing LOOP study randomizes individuals with CHA₂DS₂VASc score of ≥2 to AF screening with implantable loop recorder (ILR) screening or control.¹¹ The study's power calculation assumes that 30% of participants receiving ILR will have AF detected, and that stroke can be reduced by OAC in these patients. The primary endpoint will be published according to protocol, when the required number of events is reached.

The current substudy had two aims: First, we sought to assess incidence of AF using very long-term continuous monitoring in a large general population at risk. In this regard, we wished to validate if the AF detection rate among the earliest included participants in the LOOP study matched the study's power calculation. Second, we sought to investigate predictors of AF episodes of shorter and longer duration.

METHODS

Study design

The LOOP study is an ongoing, investigator-initiated, multicenter controlled trial. A detailed description of the study design has been published previously.¹¹ In brief, participants from the general population residing in 3 of the 5 administrative regions of Denmark are identified by administrative registries and receive a letter of invitation from one of the four study centers. Eligible subjects are ≥ 70 years old and have ≥ 1 of the following stroke risk factors; hypertension, diabetes, heart failure or previous stroke. Exclusion criteria include OAC or contraindication to OAC and any history of AF. At the initial screening visit, study eligibility is confirmed, and a baseline evaluation is performed including detailed medical history, height and weight, and blood pressure and sinus rate measurement after 10 minutes of supine rest. Prevalent AF is ruled out by 12-lead ECG. Blood is sampled for measurement of creatinine, N-terminal prohormone of brain natriuretic peptide (NT-proBNP), high-sensitivity C-reactive protein (hs-CRP), and troponin T. In the smallest center, NT-proBNP was not routinely measured for logistic reasons, and in the two smallest centers, troponin I was measured instead of troponin T. Subjects are then randomized in a 1:3 ratio to receive ILR (Reveal LINQ™, Medtronic) with continuous electrocardiographic monitoring or control. The programmable parameters of the ILRs are set to the standard for detection of "suspected AF" according to manufacturer's recommendations: AF detection "less sensitive", ectopy rejection "nominal", blank after sense 150 ms, sensing threshold decay delay 150 ms, and R-wave sensitivity

0.035 mV, though in participants with R-wave amplitude below 0.3 mV even after re-positioning, sensitivity would be re-programmed at the discretion of the implanting physician, and in participants with false AF alerts due to premature beats during follow-up, ectopy rejection would be re-programmed to “aggressive” on the discretion of the monitoring physician. All participants receiving ILR are followed by automated remote transmissions to an online database, where any new arrhythmia episodes are reviewed daily by an experienced medical doctor. AF adjudication is obtained by at least two senior cardiologists independently reviewing new-onset AF episodes lasting ≥ 6 minutes. The length of the rhythm strip used for adjudication is 2 minutes for all AF episodes. When AF is confirmed, OAC is initiated, and monitoring continues for further detection of longer AF episodes or other arrhythmias as adjudicated by minimum one experienced medical doctor. Patients are questioned about AF related symptoms at the index episode.

The heart rhythm monitoring continues until end of battery-life, study withdrawal or death, but is expected to last minimum 3 years according to the ILR manufacturer’s projected battery longevity. In the current analysis, all LOOP study participants receiving ILR until June 1st 2015 were included, and data acquisition and AF adjudication concluded on December 1st 2018. Thus, the time-span from implantation to last day of possible interrogation was minimum 42 months.

The primary endpoint of the current analysis was time to first adjudicated AF episode lasting ≥ 6 minutes, while time to first AF episode lasting ≥ 30 minutes, ≥ 1 hour, ≥ 5.5 hours, ≥ 12 hours, and ≥ 24 hours were secondary endpoints. In this way, participants reaching the endpoint of AF lasting 6 minutes at one day were subsequently followed for AF episodes of longer durations, while participants who debuted with AF lasting 24 hours at one day also reached the endpoints of AF ≥ 12 hours, ≥ 5.5 hours, ≥ 1 hour etc. on that day. Participants were censored on the date of the last ILR interrogation.

All study participants gave written informed consent. A centralized, online case report file system was used for storage of study data. The LOOP study was approved by the Ethics Committee of the

Capital Region of Denmark (H-4-2013-025) and the Danish Data Protection Agency (2007-58-0015).

The trial is registered at ClinicalTrials.gov (NCT02036450).

Statistics

Continuous variables were presented as means and standard deviations (sd) for normally distributed variables, and medians and quartiles 1 and 3 [Q1;Q3] for non-normally distributed variables, while categorical variables were presented as frequencies and corresponding percentages. Any prevalences, cumulative incidences, or rates were presented as percentage (95% confidence interval (CI)) and ratios as ratio (95% CI).

A power calculation was conducted for analyses of association between baseline variables and time to event. Assuming that 30% of the population would reach the primary end-point during follow-up, inclusion of 600 participants would yield 80% power to detect a hazard ratio (HR) of 1.5 or higher for the half of the population with the predictor (e.g. biomarker above median) compared to without the predictor (e.g. biomarker below median), with 2-sided equality and 5% risk of type I error.

Incidence rates and exact confidence intervals were derived from a Poisson distribution. To account for the competing risk of death, the cumulative incidences were estimated, plotted and group-wise compared in a multi-state fashion.¹²

To assess potential predictors of AF, association analyses were performed. Baseline variables were analyzed with univariate cause-specific Cox regression, and after considering clinical and statistical significance ($p < 0.1$) in the univariate analyses, multivariable Cox regression models were constructed to investigate the association between baseline variables and AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours independently of sex, age, and comorbidities (heart failure, hypertension, diabetes, previous stroke, cardiac valvular disease, and previous acute myocardial infarction and/or coronary arterial bypass graft surgery (CABG)), constituting a basic model. Schoenfeld and Martingale residuals were assessed to validate the proportional hazards and linearity assumption, respectively.

To comply with these assumptions for valid Cox regression, CHA₂DS₂VASc and CHADS₂ score were grouped as ≤ 3 and > 3 , and ≤ 1 and > 1 , respectively. NT-proBNP and hs-CRP were logaritmized to normalize the distribution.

Additionally, to evaluate the added discriminative value over the basic model from the physical and biochemical markers identified in the association studies as possible predictors, risk prediction analyses were conducted for the endpoints of AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours. The 42-months risk of AF was predicted using cause-specific Cox regression. Receiver operating characteristics (ROC) curves were drawn and the time-dependent area under the ROC curve (AUC) was calculated for each model. Differences in AUC between models were calculated, while Brier scores were used to assess model calibration.¹³ Furthermore, to extrapolate the predictive value of the models from the study population, the risk prediction analyses were repeated after splitting the dataset in training and test sets. The training set comprised 2/3 of the population, and analyses were repeated in 1001 splits at random. The split corresponding to the median AUC of the models without the predictors was used to analyze the difference between the basic model and the model with the predictors. All association and risk prediction was performed as complete-case analyses, meaning that missing variables in the proposed model were dropped.

Lastly, to further evaluate the relationship between continuous variables and AF episodes ≥ 6 minutes, a sliding-windows approach was used to plot the incidence rate against the median value of the variable in subgroups comprising 10% of the study population with sequentially overlapping steps the size of 2.5% of the population.¹⁴

For supplementary analyses, in the association analyses instead of adjusting for sex, age and all individual comorbidities, CHA₂DS₂VASc and CHADS₂ score, respectively, were entered into a model adjusting for each of the following baseline variables: body mass index (BMI), resting sinus rate, NT-proBNP and troponin T.

Analyses were performed using the R software, <https://www.R-project.org/>, R Core Team (2017), including the *survival*, *cmprsk*, *epiR*, *evobiR*, *timereg*, *pec*, *riskRegression*, and *ggplot2* packages.

Funding

This investigator-initiated study was supported by The Innovation Fund Denmark [12-135225], The Research Foundation for the Capital Region of Denmark [no grant number], The Danish Heart Foundation [11-04-R83-A3363-22625], Aalborg University Talent Management Programme [no grant number], Arvid Nilssons Fond [no grant number], Skibsreder Per Henriksen, R. og Hustrus Fond [no grant number], and Medtronic [no grant number]. The authors are solely responsible for the design and conduct of this study, all study analyses, the drafting and editing of the paper and its final contents.

RESULTS

Population and follow-up

A total of 597 study participants received ILR from February 26th 2014 until June 1st 2015 and were included in the current analysis. Mean age was 76 (± 4) years, 57% were men, and mean CHA₂DS₂VASc score was 3.9 (± 1.2) (Table 1). NT-proBNP was missing in 65, troponin T in 198, blood pressure in 1, resting sinus rate in 3, hs-CRP in 1, and creatinine in 2 participants.

Time from device implantation to last interrogation was median 40 [37;42] months, and 81% were monitored for minimum 3 years, yielding a total of 1897 person-years of continuous monitoring.

During this time 30 deaths occurred, and the mortality rate was 1.6 (1.1-2.2) per 100 person-years.

AF detection rate

The primary endpoint of an adjudicated AF episode lasting ≥ 6 minutes was reached in 209 participants, corresponding to a prevalence of 35% during the entire follow-up period. A total of 22

patients (11% of all with AF) reported symptoms at the index episode. The cumulative incidences at 3 years were 33.8 (30.2-37.8), 16.1 (13.4-19.4), and 5.7 (4.1-7.9)% for AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours and ≥ 24 hours, respectively (Figure 1).

The incidence rate of AF ≥ 6 minutes was 14.5 (12.7-16.6) per 100 person-years, and most incident cases of short episodes were detected during the first year after implantation, while the incidence rate was more stable for longer AF episodes (Supplementary Figure 1). Time to detection was median 5.5 [1.4;17], 14 [4.8;26], and 20 [6.4;30] months for the first AF episode lasting ≥ 6 minutes, ≥ 5.5 hours and ≥ 24 hours, respectively (Figure 2).

Association analyses and risk prediction

Table 1 shows baseline characteristics of the study population according to AF detection. Univariate Cox regressions of association between baseline variables and time to AF detection are shown in Supplementary Table 2, and cumulative incidence of AF of different durations by groups of baseline markers are shown in Supplementary Figure 2. After multivariate adjustment higher age, NT-proBNP, troponin T, hs-CRP, and slower resting sinus rate were independently associated with detection of time to AF ≥ 6 minutes (Figure 3). For longer episodes, also higher BMI and previous CABG were associated with AF, whereas age and hs-CRP did not reach statistical significance.

Supplementary Table 1 presents baseline characteristics according to whether any markers required for risk prediction analyses were missing. There were no relevant differences among patients included in risk prediction analyses or not. In the risk prediction analyses, the basic model of sex, age, and comorbidities had poor discriminative value for screen-detected AF, while addition of baseline BMI, resting sinus rate, NT-proBNP and troponin T significantly increased AUC for AF episodes ≥ 24 hours, 79% vs. 65%, $p=0.037$ (Figure 4). This was also true after the data was split in training and test sets (Supplementary Table 3). For all models, Brier score was lower when the markers were included.

The sliding-windows approach showed decreased incidence rate of AF episodes ≥ 24 hours in participants with BMI $< 26 \text{ kg/m}^2$, resting sinus rate > 62 beats per minute, NT-proBNP $< 12 \text{ pmol/L}$, troponin T $< 12 \text{ ng/L}$ (Figure 5). Comparable results were seen for shorter episodes, though not as discernible.

In the supplementary Cox regression models adjusted for markers of BMI, resting sinus rate, NT-proBNP, and troponin T, CHA₂DS₂VASc score group was not associated with AF of any duration, and CHADS₂ score > 1 was only associated with AF ≥ 6 minutes (HR 1.59 (1.02-2.5)).

DISCUSSION

The present study is the first to investigate AF in a large general population at risk using long-term monitoring corresponding to the lifetime of a modern ILR device. First, we found that AF was detected in 35% of participants, which is in agreement with the assumptions of the LOOP study protocol. Second, approximately 16% of all participants had episodes exceeding the debatable duration-threshold for OAC in device-detected AF,⁹ and 50% of these episodes were detected after more than 14 months of monitoring. Third, slower resting sinus rate and higher BMI, NT-proBNP, and troponin T at baseline were independently associated with AF detection, and addition of these markers to a model of sex, age, and comorbidities significantly improved prediction of AF episodes ≥ 24 hours.

ILR screening studies

The low frequency of symptoms at AF debut (11%) underlines that we have indeed investigated a subclinical arrhythmia. The detection rates in the current study can be compared to findings in recent ILR screening studies.^{3,14-16} However, these studies were not powered to investigate predictors of AF. The CRYSTAL-AF study monitored 221 patients with recent cryptogenic stroke or transient ischemic attack and found AF in 12% at 12 months.³ The ASSERT-II study recruited 256 patients with risk

factors from cardiology and neurology clinics and monitored them during approximately 12 months in which 31% reached the endpoint of AF episodes lasting ≥ 5 minutes.¹⁴ Age and left atrial diameter were associated with time to AF, while higher blood pressure was inversely associated with AF detection, possibly indicating selection bias. The PREDATE AF study found AF ≥ 6 minutes in 55 of 245 patients (22%) during approximately 15 months,¹⁵ while the REVEAL-AF study monitored 326 patients for at least 18 months at which an estimated cumulative rate of 29% was reported.¹⁶

The lower rate of AF detection in CRYSTAL-AF compared to ASSERT-II, PREDATE AF, REVEAL-AF, and the current study, respectively, could be due to lower age; mean 61.5 years compared to 73.9, 74.3, 71.6, and 76.4 years, respectively. Furthermore, all patients in CRYSTAL-AF were monitored with a device with inferior sensitivity compared to the devices used in most patients in the other studies, increasing the risk of false-negatives.¹⁷⁻¹⁹

While the participants included in the LOOP study and the ASSERT-II, PREDATE AF, and REVEAL-AF were rather similar in terms of conventional risk factors (CHA₂DS₂VASc score was mean 4.1, 4.6, and 4.4 in ASSERT-II, PREDATE AF, and REVEAL-AF, respectively), the current study is unique in having recruited participants from home as opposed to a clinical setting, which arguably decreases selection bias. Secondly, the current study provides a larger sample size and longer follow-up, yielding enough person-years to investigate AF predictors. Keeping in mind that detection of longer episodes required very long-term monitoring, the current study enables investigation of AF burden of different durations opposed to short AF episodes in general. The outcomes of AF episode duration were chosen to reflect the controversies in the current understanding of duration-threshold for increased stroke risk in subclinical AF; “any” AF (≥ 6 minutes),⁶ the duration in the European Heart Rhythm Association consensus statement (≥ 5.5 hours),⁹ or the duration that upheld the increased risk in ASSERT (≥ 24 hours).⁸

Predictors of AF episodes of different durations

To the best of our knowledge, the current study is the first to demonstrate an association between troponin and screen-detected AF, while a recent study inferred that NT-proBNP and BMI are useful in selection of patients for short-term screening for prevalent AF.²⁰ Both NT-proBNP, troponin, and BMI are established markers in clinically diagnosed AF.^{21–23} The inverse association between incident AF and resting sinus rate at baseline in our population can be compared to epidemiological findings of higher risk of clinical AF in both slower and faster sinus rates.²⁴ When we modelled resting sinus rate as a categorical variable, participants in the fastest tertile still had significantly lower risk than those with slowest rates, and this was independent of BMI as well as other risk factors. The sliding-windows approach confirmed the inverse relationship (Figure 5). Possible explanations for this finding could be good physical condition, or subclinical sinus node or AV conduction dysfunction leading to relative bradycardia and increased atrial preload in these patients.^{25,26} Furthermore, in this elderly population, we did not find an association between height and screen-detected AF, oppositely to what is known from clinical AF.²⁷

Of demographic variables, sex was not a risk factor for AF in our study, which is in agreement with recently evaluated risk models for clinical AF.²⁸ Furthermore, we did find an association between age and AF, though not as strong as what is described in clinical AF.²⁸ This could be due to the fact that we only included participants ≥ 70 years of age.

Interestingly, previous CABG was associated with AF episodes lasting ≥ 24 hours, as 7 of 41 patients (17%) with previous CABG had such episodes during follow-up (Table 1). However, most of the CHA₂DS₂VASc components and medical treatments did not show association with AF (Figure 3 and Supplementary Table 2). Likewise, in agreement with the above described ILR studies, higher CHA₂DS₂VASc score did not by itself predict screen-detected AF. This could be due to the fact that we only included participants with CHA₂DS₂VASc score ≥ 2 , and thus, a possible risk attributed to one CHA₂DS₂VASc component could be clouded by presence of other risk-attributes in participants

without that component. Arguably, instead of predicting AF from traditional risk factors among patients with increased risk of stroke, an important goal would be to identify easily available biomarkers to accompany the traditional risk factors in selection of patients for screening. We performed risk prediction analyses to evaluate whether the physical and biochemical markers from the association studies improved prediction over a model considering only sex, age, and previous diagnoses. These analyses revealed that this basic model was indeed poor at identifying participants at risk of AF, while addition of NT-proBNP, troponin, BMI, and resting sinus rate improved prediction for episodes lasting ≥ 24 hours. This holds important implications if the ongoing trials^{11,29,30} find that very low AF burden does not merit OAC, and future screening should thus aim at identifying patients with longer AF episodes. Indeed, a relationship between AF burden and stroke risk has been suggested.⁸ Our study shows that a consideration of biomarkers could reduce the number needed to screen.

In summary, we have presented findings from a large general population of AF-naïve participants with risk factors for stroke screened during the full life-span of a standard ILR. The data suggests that a considerable AF burden is detected when long-term continuous monitoring is applied, and that physical and biochemical markers can improve prediction of these AF episodes.

Limitations

The current study has several limitations. First, since the LOOP study invited subjects by letter, our results could possibly be affected by healthy user bias, meaning that persons with better health status are more inclined to participate. However, the rates of AF we found were comparable to those in studies of patients with cardiac implantable electronic devices.⁵ Second, the ILR device is capable of monitoring for minimum 3 years, and thus all participants receiving ILR in the LOOP study at least 3.5 years before December 2018 were included. This provided a long follow-up duration and sufficient sample size to assess predictors, although increased statistical power would likely have

revealed more variables associated with AF of different durations, and would enable investigation of which individual markers would improve AUC for AF in risk prediction analyses, as opposed to a set of markers. Furthermore, a subset of participants had missing NT-proBNP and Troponin T due to logistic reasons, and analyses including these variables have limited power. In spite of this, the biomarkers did in fact constitute a significant addition to the discriminative value of a basic model in identifying patients with increased risk of long AF episodes. Further studies are needed to confirm biomarker cutoffs in selection of patients to screening.

Our findings rely on the capability of the ILR to detect AF. The algorithm used has been shown to have a sensitivity of approximately 95%, although specificity might be decreased, especially for short episodes.^{17,31} Thus, to avoid false positives, a rigorous adjudication regimen was applied. Finally, the current results are limited to a discussion about AF detection and predictors of AF in at-risk patients recruited from the general population, and we cannot yet report on the health benefits of the screening. We anticipate further investigation when the LOOP study has been finalized, including studies of additional risk markers from magnetic resonance imaging for atrial fibrosis and function, echocardiographic measures, genetics and more, along with the main studies of impact on stroke risk and cost-effectiveness in ILR screening for AF.¹¹

Conclusion

A considerable burden of previously unknown AF was detected when long-term monitoring was applied in at-risk patients from the general population. Sex, age, and comorbidities had poor discriminative value in prediction of AF in these patients. Markers of BMI, resting sinus rate, NT-proBNP, and troponin T were associated with AF incidence and improved prediction of long AF episodes. These findings may hold important implications for future screening programs.

ACKNOWLEDGEMENTS

We thank Professor Dan Atar, University of Oslo and Department of Cardiology B, Oslo University Hospital Ullevål, Norway, Professor Gregory Y. H. Lip, The University of Liverpool and Liverpool Heart and Chest Hospital, United Kingdom, and Professor Mårten Rosenqvist, Karolinska Institutet and Danderyd Hospital, Sweden, for assisting the study with their expertise in the study's International Advisory Committee.

We thank research nurses and other colleagues in the departments of cardiology, Rigshospitalet, Zealand University Hospital Roskilde, Bispebjerg Hospital, and Odense University Hospital having assisted with patient inclusion, data collection and ILR procedures.

CONFLICTS OF INTEREST

K.J.H. reports travel and educational grants from Medtronic, Boston Scientific, Abbott, and Biotronik not related to this work. A.B. reports a research grant from Gilead and personal fees from Bayer, Boehringer Ingelheim, MSD, and Bristol-Myers Squibb not related to this work. D.K. reports to be a Medtronic Focus Group member. A.G.H. is an employee of Novo Nordisk, not related to this work. L.K. reports speaker honoraria from, Bayer, Astra-Zeneca, Orion Pharma, Novartis, and Sanofi, not related to this work. J.H.S. reports to be a member of Medtronic advisory boards and to have received speaker honoraria and research grants from Medtronic in relation to this work, in addition to research grant from Gilead, research grant and personal fees from Biotronik and personal fees from Boehringer Ingelheim, Novo Nordisk, Bayer, and Astra-Zeneca not related to this work.

REFERENCES

1. Heidenreich PA, Trogdon JG, Khavjou OA, Butler J, Dracup K, Ezekowitz MD, *et al.* Forecasting the future of cardiovascular disease in the United States: a policy statement from the American Heart Association. *Circulation* 2011;**123**:933–44.
2. Marini C, De Santis F, Sacco S, Russo T, Olivieri L, Totaro R, *et al.* Contribution of atrial fibrillation to incidence and outcome of ischemic stroke: results from a population-based study. *Stroke J Cereb Circ* 2005;**36**:1115–9.
3. Brachmann J, Morillo CA, Sanna T, Di Lazzaro V, Diener H-C, Bernstein RA, *et al.* Uncovering Atrial Fibrillation Beyond Short-Term Monitoring in Cryptogenic Stroke Patients: Three-Year Results From the Cryptogenic Stroke and Underlying Atrial Fibrillation Trial. *Circ Arrhythm Electrophysiol* 2016;**9**:e003333.
4. Kirchhof P, Benussi S, Kotecha D, Ahlsson A, Atar D, Casadei B, *et al.* 2016 ESC Guidelines for the management of atrial fibrillation developed in collaboration with EACTS. *Eur Heart J* 2016;**37**:2893–962.
5. Healey JS, Connolly SJ, Gold MR, Israel CW, Van Gelder IC, Capucci A, *et al.* Subclinical atrial fibrillation and the risk of stroke. *N Engl J Med* 2012;**366**:120–9.
6. Mahajan R, Perera T, Elliott AD, Twomey DJ, Kumar S, Munwar DA, *et al.* Subclinical device-detected atrial fibrillation and stroke risk: a systematic review and meta-analysis. *Eur Heart J* 2018;**39**:1407–15.
7. Mairesse GH, Moran P, Van Gelder IC, Elsner C, Rosenqvist M, Mant J, *et al.* Screening for atrial fibrillation: a European Heart Rhythm Association (EHRA) consensus document endorsed by the Heart Rhythm Society (HRS), Asia Pacific Heart Rhythm Society (APHRS), and Sociedad Latinoamericana de Estimulación Cardíaca y Electrofisiología (SOLAECE). *EP Eur* 2017;**19**:1589–623.
8. Van Gelder I., Healey JS, Crijns HJGM, Wang J, Hohnloser SH, Gold MR, *et al.* Duration of device-detected subclinical atrial fibrillation and occurrence of stroke in ASSERT. *Eur Heart J* 2017;**38**:1339–44.
9. Gorenek B, Bax J, Boriani G, Chen S-A, Dagres N, Glotzer TV, *et al.* Device-detected subclinical atrial tachyarrhythmias: definition, implications and management—an European Heart Rhythm Association (EHRA) consensus document, endorsed by Heart Rhythm Society (HRS), Asia Pacific Heart Rhythm Society (APHRS) and Sociedad Latinoamericana de Estimulación Cardíaca y Electrofisiología (SOLEACE). *Europace* 2017;**19**:1556–78.
10. Perino Alexander C., Fan Jun, Askari Mariam, Heidenreich Paul A., Keung Edmund, Raitt Merritt H., *et al.* Practice Variation in Anticoagulation Prescription and Outcomes After Device-Detected Atrial Fibrillation. *Circulation* 2019;**139**:2502–12.
11. Diederichsen SZ, Haugan KJ, Køber L, Højberg S, Brandes A, Kronborg C, *et al.* Atrial fibrillation detected by continuous electrocardiographic monitoring using implantable loop recorder to prevent stroke in individuals at risk (the LOOP study): Rationale and design of a large randomized controlled trial. *Am Heart J* 2017;**187**:122–32.

12. Aalen O. Nonparametric Estimation of Partial Transition Probabilities in Multiple Decrement Models. *Ann Stat* 1978;**6**:534–45.
13. Blanche P, Proust-Lima C, Loubère L, Berr C, Dartigues J-F, Jacqmin-Gadda H. Quantifying and comparing dynamic predictive accuracy of joint models for longitudinal marker and time-to-event in presence of censoring and competing risks. *Biometrics* 2015;**71**:102–13.
14. Healey JS, Alings M, Ha A, Leong-Sit P, Birnie DH, Graaf JJ de, *et al.* Subclinical Atrial Fibrillation in Older Patients. *Circulation* 2017;**136**:1276–83.
15. Nasir JM, Pomeroy W, Marler A, Hann M, Baykaner T, Jones R, *et al.* Predicting Determinants of Atrial Fibrillation or Flutter for Therapy Elucidation in Patients at Risk for Thromboembolic Events (PREDATE AF) Study. *Heart Rhythm* 2017;**14**:955–61.
16. Reiffel JA, Verma A, Kowey PR, Halperin JL, Gersh BJ, Wachter R, *et al.* Incidence of Previously Undiagnosed Atrial Fibrillation Using Insertable Cardiac Monitors in a High-Risk Population. *JAMA Cardiol* 2017;**2**:1120–7.
17. Mittal S, Rogers J, Sarkar S, Koehler J, Warman EN, Tomson TT, *et al.* Real-world performance of an enhanced atrial fibrillation detection algorithm in an insertable cardiac monitor. *Heart Rhythm* 2016;**13**:1624–30.
18. Nölker G, Mayer J, Boldt L-H, Seidl K, VAN Driel V, Massa T, *et al.* Performance of an Implantable Cardiac Monitor to Detect Atrial Fibrillation: Results of the DETECT AF Study. *J Cardiovasc Electrophysiol* 2016;**27**:1403–10.
19. Podd SJ, Sugihara C, Furniss SS, Sulke N. Are implantable cardiac monitors the ‘gold standard’ for atrial fibrillation detection? A prospective randomized trial comparing atrial fibrillation monitoring using implantable cardiac monitors and DDDR permanent pacemakers in post atrial fibrillation ablation patients. *Eur Eur Pacing Arrhythm Card Electrophysiol J Work Groups Card Pacing Arrhythm Card Cell Electrophysiol Eur Soc Cardiol* 2016;**18**:1000–5.
20. Chua W, Purmah Y, Cardoso VR, Gkoutos GV, Tull SP, Neculau G, *et al.* Data-driven discovery and validation of circulating blood-based biomarkers associated with prevalent atrial fibrillation. *Eur Heart J*.
21. Filion KB, Agarwal SK, Ballantyne CM, Eberg M, Hoogeveen RC, Huxley RR, *et al.* High-sensitivity cardiac troponin T and the risk of incident atrial fibrillation: The Atherosclerosis Risk in Communities (ARIC) study. *Am Heart J* 2015;**169**:31–38.e3.
22. Svennberg E, Lindahl B, Berglund L, Eggers KM, Venge P, Zethelius B, *et al.* NT-proBNP is a powerful predictor for incident atrial fibrillation — Validation of a multimarker approach. *Int J Cardiol* 2016;**223**:74–81.
23. Frost L, Hune LJ, Vestergaard P. Overweight and obesity as risk factors for atrial fibrillation or flutter: The Danish Diet, Cancer, and Health Study. *Am J Med* 2005;**118**:489–95.
24. Skov MW, Bachmann TN, Rasmussen PV, Olesen MS, Pietersen A, Graff C, *et al.* Association Between Heart Rate at Rest and Incident Atrial Fibrillation (from the Copenhagen Electrocardiographic Study). *Am J Cardiol* 2016;**118**:708–13.

25. Aladin AI, Al Rifai M, Rasool SH, Keteyian SJ, Brawner CA, Blumenthal RS, *et al.* Relation of Resting Heart Rate to Incident Atrial Fibrillation (from the Henry Ford Hospital Exercise Testing Project). *Am J Cardiol* 2017;**119**:262–7.
26. Nielsen JC, Thomsen PEB, Højberg S, Møller M, Riahi S, Dalsgaard D, *et al.* Atrial fibrillation in patients with sick sinus syndrome: the association with PQ-interval and percentage of ventricular pacing. *Europace* 2012;**14**:682–9.
27. Rosenberg MA, Patton KK, Sotoodehnia N, Karas MG, Kizer JR, Zimetbaum PJ, *et al.* The impact of height on the risk of atrial fibrillation: the Cardiovascular Health Study. *Eur Heart J* 2012;**33**:2709–17.
28. Linker DT, Murphy TB, Mokdad AH. Selective screening for atrial fibrillation using multivariable risk models. *Heart* 2018;**104**:1492–9.
29. Lopes RD, Alings M, Connolly SJ, Beresh H, Granger CB, Mazuecos JB, *et al.* Rationale and design of the Apixaban for the Reduction of Thrombo-Embolism in Patients With Device-Detected Sub-Clinical Atrial Fibrillation (ARTESiA) trial. *Am Heart J* 2017;**189**:137–45.
30. Kirchhof P, Blank BF, Calvert M, Camm AJ, Chlouverakis G, Diener H-C, *et al.* Probing oral anticoagulation in patients with atrial high rate episodes: Rationale and design of the Non-vitamin K antagonist Oral anticoagulants in patients with Atrial High rate episodes (NOAH-AFNET 6) trial. *Am Heart J* 2017;**190**:12–8.
31. Sanders P, Pürerfellner H, Pokushalov E, Sarkar S, Di Bacco M, Maus B, *et al.* Performance of a new atrial fibrillation detection algorithm in a miniaturized insertable cardiac monitor: Results from the Reveal LINQ Usability Study. *Heart Rhythm* 2016;**13**:1425–30.

FIGURE LEGENDS

Figure 1

Title: Cumulative incidence of AF episodes

Legend: The figure presents cumulative incidence of AF detection plotted with the Aalen-Johansen method.

Figure 2

Title: Time to detection of AF episodes

Legend: The figure presents boxplots of months from device implantation to AF detection. The vertical lines represent quartile 1, median and quartile 3, while the horizontal line represents the range and the diamond represents the mean.

Figure 3

Title: Association between baseline variables and time to AF after multivariate adjustment

Legend: The figure represents hazard ratios and confidence intervals for the association between baseline markers and time to AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours in cause-specific Cox models adjusting for sex, age, and comorbidities (heart failure, hypertension, diabetes, previous stroke, cardiac valvular disease, and previous acute myocardial infarction and/or CABG). The models of age, previous stroke, previous CABG, and BMI included 597 patients, while the models of resting sinus rate, NT-proBNP, troponin T and hs-CRP included 594, 532, 399, and 596 patients, respectively. The CABG model did not include previous acute myocardial infarction. The previous stroke model was unchanged when stroke and/or transient ischemic attack was included instead of stroke alone. The resting sinus rate model further included body mass index, which did not change the estimates, and the results were similar when resting sinus rate was modelled as a categorical variable of tertiles. Finally, the estimates for AF ≥ 6 minutes and 5.5 hours were unchanged when the models of resting sinus rate were applied only to patients not treated with betablockers or cardiac calcium blockers ($n=440$), while for AF ≥ 24 hours significance was not reached in this strata.

Figure 4

Title: Prediction of AF using the basic model or the model including additional markers

Legend: The figure presents time-dependent receiver operating curves for two models in prediction of AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours along with P-values for difference. The *Basic model* is a cause-specific Cox model including sex, age, and comorbidities, while the *Basic model + markers* further included N-terminal pro-brain natriuretic peptide, troponin T, body mass index and resting sinus rate. These analyses were performed in exactly the same population, and due to missing blood biomarker information, the sample comprised a total of 397 patients.

Figure 5

Title: Relationship between baseline markers and incidence rate of AF

Legend: The y-axis presents exact %/year incidence rate (points) and confidence intervals (colored ribbon) for AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours in sequentially overlapping subgroups comprising 10% of the population on the x-axis. The curve represents a moving average using local polynomial regression fitting, while the dashed line represents the average incidence rate for the entire population.

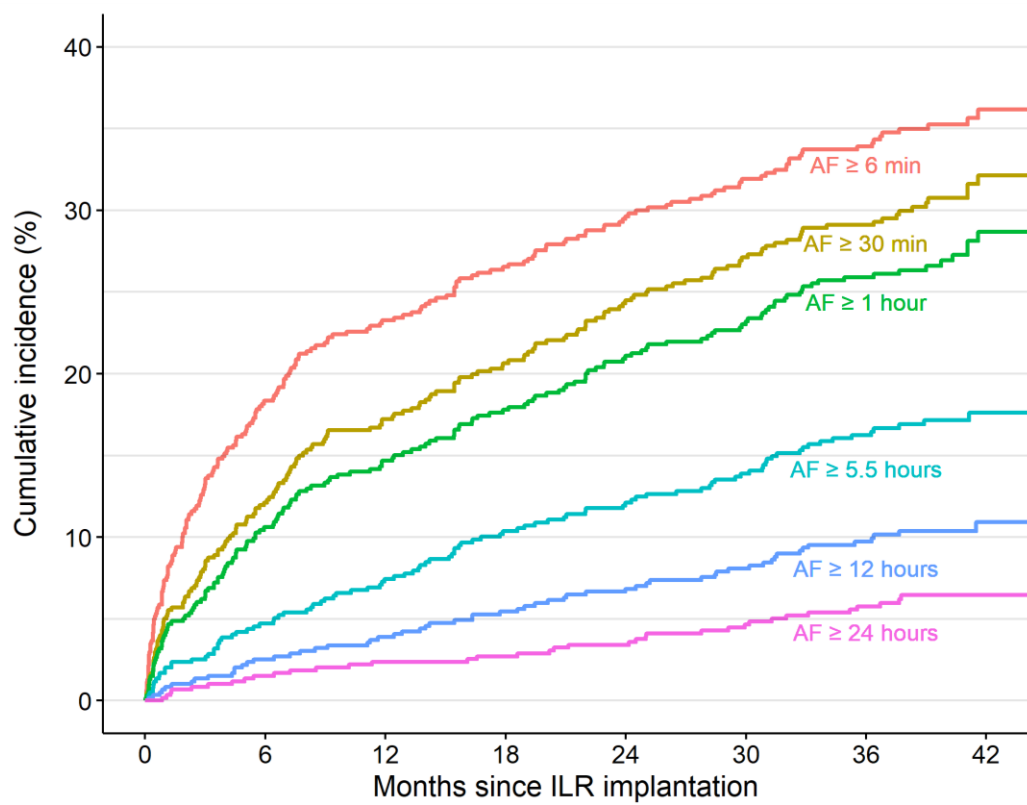
Table 1

Baseline characteristics for patients without AF and AF episodes of different durations

AF during follow-up n (% of all)	No AF 388 (65)	AF ≥ 6 min 209 (35)	AF ≥ 30 min 180 (30)	AF ≥ 1 hour 159 (27)	AF ≥ 5.5 hours 99 (17)	AF ≥ 12 hours 61 (10)	AF ≥ 24 hours 37 (6)	All 597 (100)
Male sex (%)	218 (56)	123 (59)	105 (58)	91 (57)	63 (64)	42 (69)	25 (68)	341 (57.1)
Age, years (sd)	76.0 (4.0)	77.1 (4.5)	77.0 (4.4)	77.1 (4.6)	77.1 (4.7)	76.8 (4.4)	77.2 (4.6)	76.4 (4.2)
CHA ₂ DS ₂ VASc score (sd)	3.9 (1.2)	4.1 (1.2)	4.0 (1.2)	4.1 (1.2)	4.1 (1.2)	4.0 (1.1)	4.0 (1.1)	3.9 (1.2)
CHADS ₂ score (sd)	2.2 (1.1)	2.4 (1.1)	2.4 (1.1)	2.4 (1.0)	2.5 (1.0)	2.4 (1.0)	2.4 (1.0)	2.3 (1.1)
Heart failure (%)	19 (4.9)	6 (2.9)	5 (2.8)	4 (2.5)	2 (2.0)	2 (3.3)	1 (2.7)	25 (4.2)
Hypertension (%)	350 (90)	190 (91)	165 (92)	146 (92)	90 (91)	56 (92)	34 (92)	540 (90.5)
Diabetes (%)	114 (29)	60 (29)	51 (28)	46 (29)	30 (30)	18 (30)	11 (30)	174 (29.1)
Previous stroke (%)	64 (16)	43 (21)	37 (21)	34 (21)	23 (23)	16 (26)	9 (24)	107 (17.9)
Previous transient ischemic attack, n (%)	46 (11.9)	21 (10.0)	19 (10.6)	16 (10.1)	10 (10.1)	5 (8.2)	3 (8.1)	67 (11.2)
Previous systemic embolism, n (%)	25 (6.4)	17 (8.1)	15 (8.3)	14 (8.8)	10 (10.1)	5 (8.2)	2 (5.4)	42 (7.0)
Previous AMI (%)	38 (9.8)	20 (9.6)	17 (9.4)	16 (10.1)	12 (12.1)	8 (13.1)	5 (13.5)	58 (9.7)
Previous CABG (%)	27 (7.0)	14 (6.7)	14 (7.8)	13 (8.2)	11 (11.1)	9 (14.8)	7 (18.9)	41 (6.9)
Valvular heart disease (%)	13 (3.3)	13 (6.2)	8 (4.4)	6 (3.8)	4 (4.0)	2 (3.3)	1 (2.7)	26 (4.4)
Beta blockers, n (%)	95 (24)	49 (23)	45 (25.0)	41 (25.8)	29 (29.3)	23 (37.7)	15 (40.5)	144 (24.1)
Calcium channel blockers, n (%)	124 (32)	85 (41)	72 (40.0)	63 (39.6)	43 (43.4)	32 (52.5)	17 (45.9)	209 (35.0)
Non-dihydropyridine type, n (%)	5 (1.3)	9 (4.3)	8 (4.4)	6 (3.8)	3 (3.0)	2 (3.3)	0 (0.0)	14 (2.3)
ACEi, ARB, or renin inhibitors, n (%)	235 (61)	121 (58)	105 (58.3)	93 (58.5)	57 (57.6)	35 (57.4)	24 (64.9)	356 (59.6)
Lipid-lowering drugs, n (%)	201 (52)	121 (58)	107 (59.4)	97 (61.0)	65 (65.7)	39 (63.9)	23 (62.2)	322 (53.9)
Diuretics, n (%)	112 (29)	65 (31)	57 (31.7)	52 (32.7)	34 (34.3)	19 (31.1)	12 (32.4)	177 (29.6)
Platelet inhibitors, n (%)	188 (48)	107 (51)	93 (51.7)	84 (52.8)	54 (54.5)	36 (59.0)	23 (62.2)	295 (49.4)
Glucose-lowering drugs, n (%)	98 (25)	51 (24)	45 (25.0)	40 (25.2)	27 (27.3)	17 (27.9)	11 (29.7)	149 (25.0)
Systolic BP, mmHg (sd)	151.9 (19)	151.6 (18)	151.7 (18)	152.2 (18)	151.9 (17)	149.6 (18)	153.4 (17)	151.8 (18.7)
Diastolic BP, mmHg (sd)	85.0 (12)	84.6 (11)	84.7 (11)	84.6 (11)	84.5 (11)	84.1 (11)	86.0 (12)	84.9 (11.6)
Resting sinus rate, bpm (sd)	72.5 (13)	68.8 (11)	68.7 (12)	68.6 (12)	68.0 (12)	68.2 (13)	67.1 (13)	71.2 (12.4)
Height, cm (sd)	170.5 (8.7)	171.3 (8.7)	171.2 (8.8)	171.0 (8.8)	171.6 (8.5)	172.9 (8.6)	172.9 (9.0)	170.8 (8.7)
Body mass index, kg/m ² (sd)	27.4 (4.4)	27.8 (4.8)	27.6 (4.5)	27.8 (4.6)	28.3 (4.8)	28.4 (4.7)	29.3 (4.9)	27.6 (4.6)
Creatinine, μmol/L (sd)	86.7 (23)	87.9 (24)	88.3 (25)	87.9 (24)	89.8 (24)	92.1 (25)	91.7 (24)	87.1 (23.6)
NT-proBNP, pmol/L [Q1;Q3]	14 [8;26]	20 [12;34]	20 [12;34]	21 [12;34]	21 [12;34]	24 [13;38]	24 [13;44]	16 [9;28]
hs-CRP, mg/L [Q1;Q3]	2 [1;3]	2 [1;4]	2 [1;4]	2 [1;4]	2 [1;4]	2 [1;4]	2 [1;4]	2 [1;4]
Troponin T, ng/L (sd)	13.8 (5.4)	15.8 (7.3)	15.8 (7.9)	15.9 (8.3)	16.4 (8.9)	17.5 (10.4)	19.0 (12.1)	14.6 (6.48)

The table presents baseline characteristics of the study participants according to presence of AF.

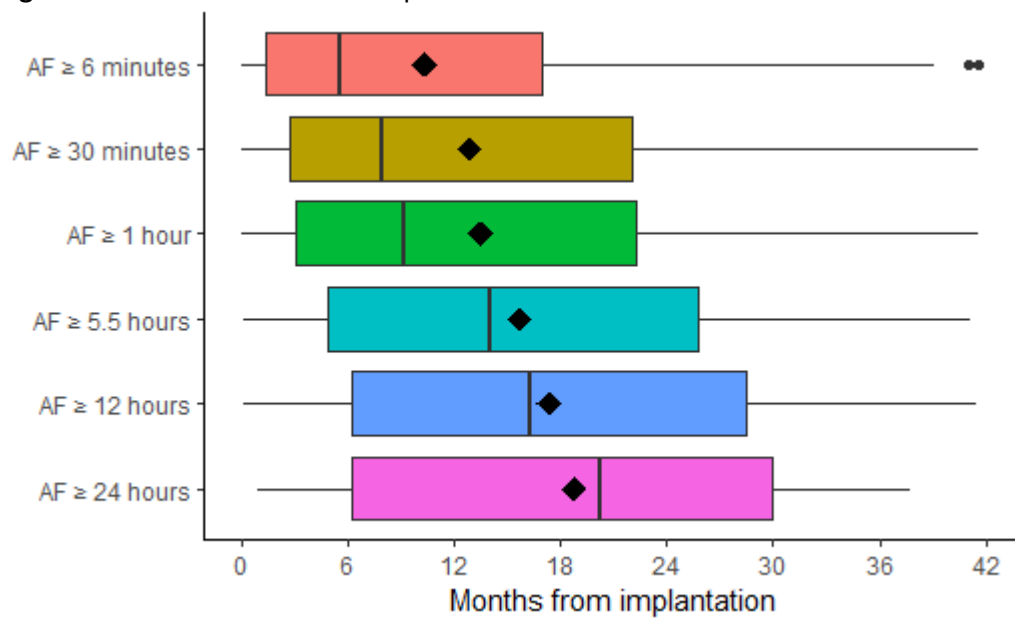
Abbreviations: AF, atrial fibrillation; AMI, acute myocardial infarction; CABG, coronary artery bypass graft surgery; NT-proBNP, N-terminal prohormone of brain natriuretic peptide; hs-CRP, high-sensitivity C-reactive protein.

Figure 1 Cumulative incidence of AF episodes

AF duration	No. at risk							
≥ 6 min	597	483	449	427	404	382	326	107
≥ 30 min	597	519	484	460	432	408	348	110
≥ 1 hour	597	528	498	476	451	430	363	115
≥ 5.5 hours	597	562	540	518	501	479	409	129
≥ 12 hours	597	575	561	544	529	510	437	135
≥ 24 hours	597	581	570	560	549	530	459	138

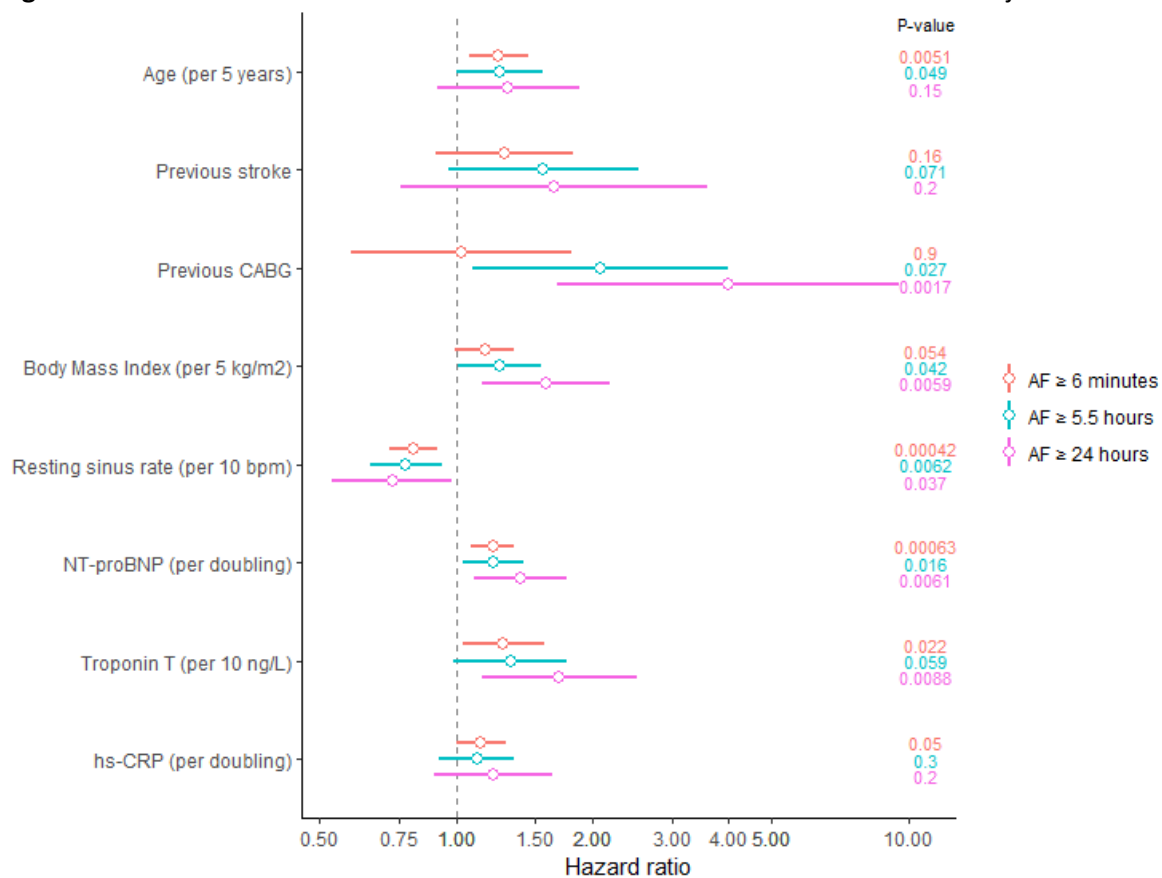
Legend: The figure presents cumulative incidence of AF detection plotted with the Aalen-Johansen method.

Abbreviations: AF, atrial fibrillation; ILR, implantable loop recorder; min, minutes.

Figure 2 Time to detection of AF episodes

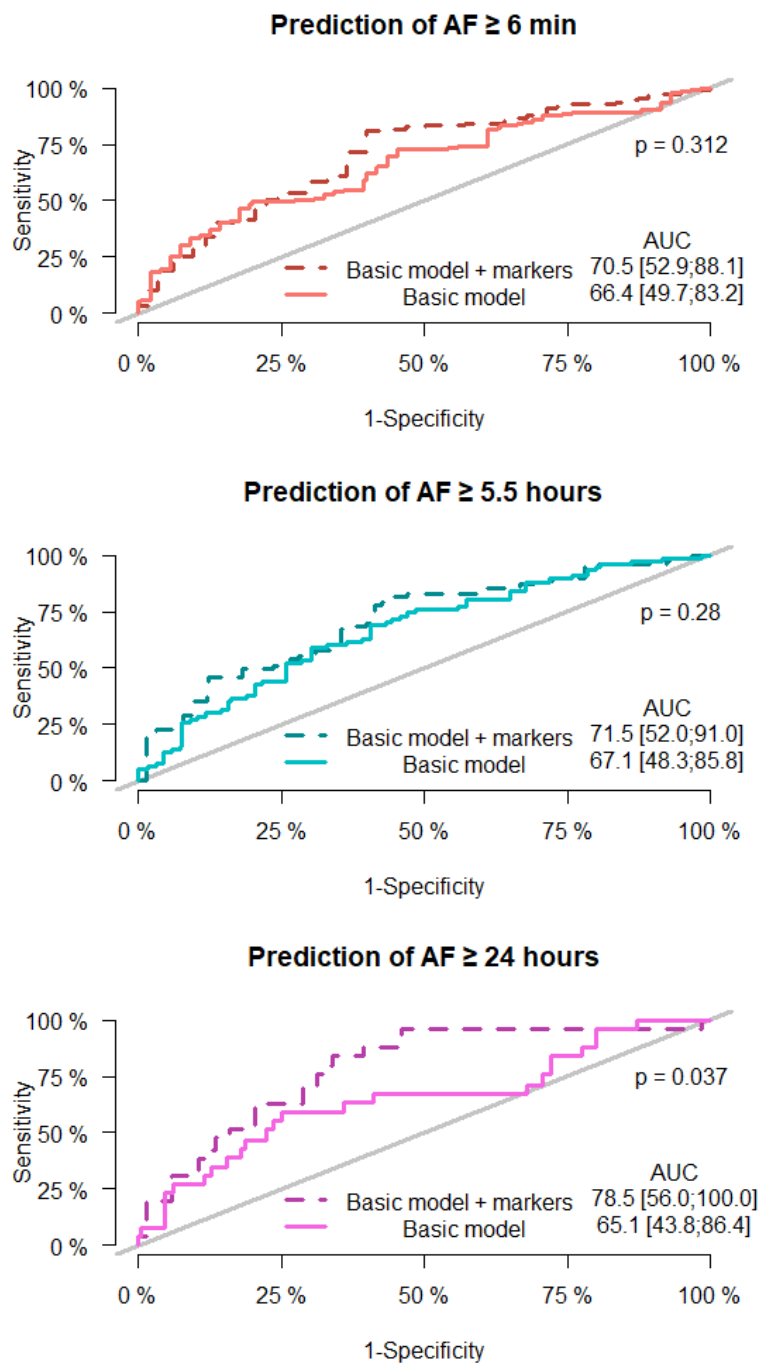
Legend: The figure presents boxplots of months from device implantation to AF detection. The vertical lines represent quartile 1, median and quartile 3, while the horizontal line represents the range and the diamond represents the mean.

Abbreviations: AF, atrial fibrillation.

Figure 3 Association between baseline variables and time to AF after multivariate adjustment

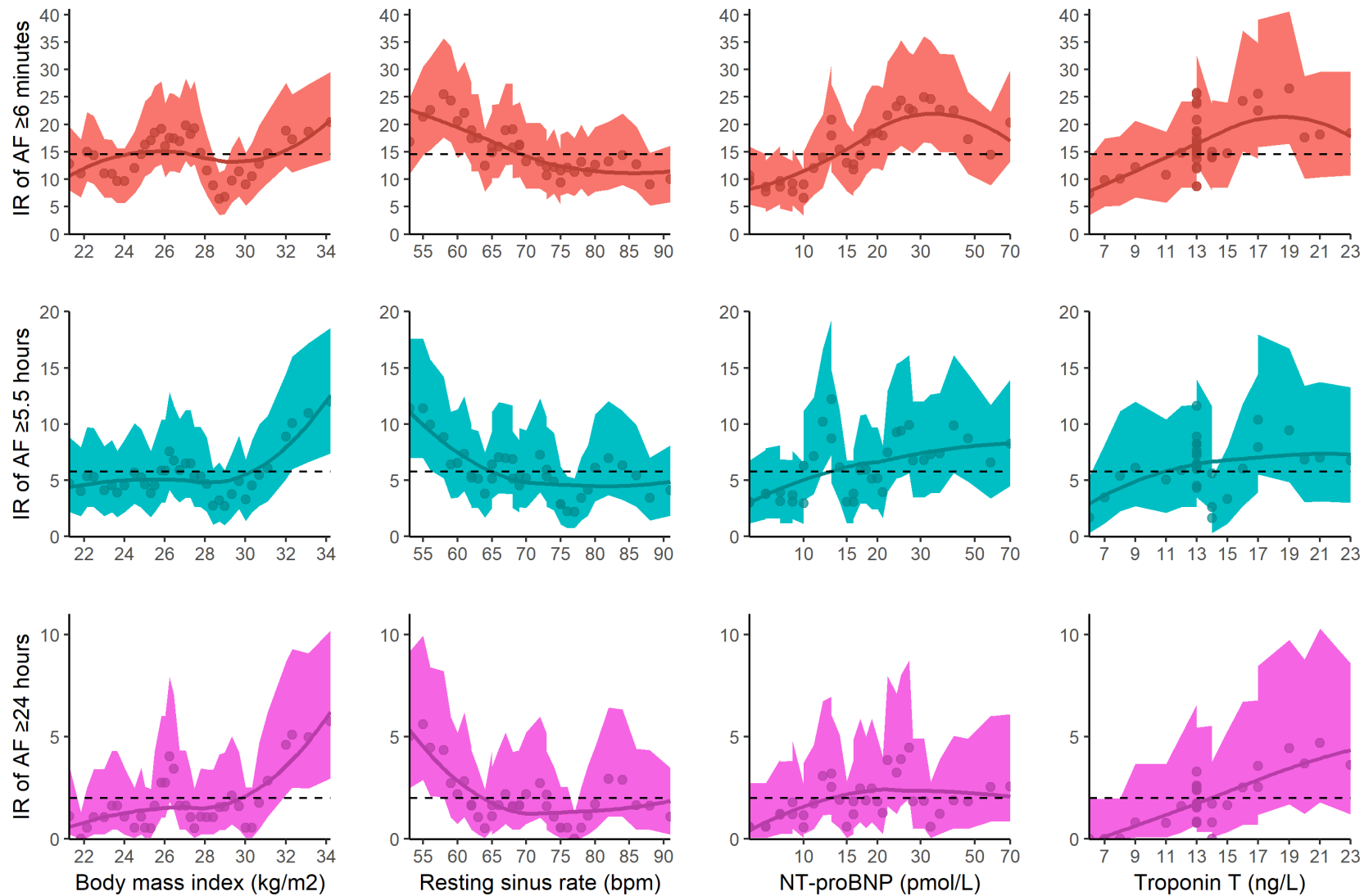
Legend: The figure represents hazard ratios and confidence intervals for the association between baseline markers and time to AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours in cause-specific Cox models adjusting for sex, age, and comorbidities (heart failure, hypertension, diabetes, previous stroke, cardiac valvular disease, and previous acute myocardial infarction and/or CABG). The models of age, previous stroke, previous CABG, and BMI included 597 patients, while the models of resting sinus rate, NT-proBNP, troponin T and hs-CRP included 594, 532, 399, and 596 patients, respectively. The CABG model did not include previous acute myocardial infarction. The previous stroke model was unchanged when stroke and/or transient ischemic attack was included instead of stroke alone. The resting sinus rate model further included body mass index, which did not change the estimates, and the results were similar when resting sinus rate was modelled as a categorical variable of tertiles. Finally, the estimates for AF ≥ 6 minutes and ≥ 5.5 hours were unchanged when the models of resting sinus rate were applied only to patients not treated with betablockers or non-dihydropyridine calcium blockers ($n=440$), while significance was not reached in this strata for AF ≥ 24 hours.

Abbreviations: AF, atrial fibrillation; CABG, coronary artery bypass graft surgery; hs-CRP, high-sensitivity C-reactive protein; NT-proBNP, N-terminal prohormone of brain natriuretic peptide.

Figure 4 Prediction of AF using the basic model or the model including additional markers

Legend: The figure presents time-dependent receiver operating curves for two models in prediction of AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours along with P-values for difference. The *Basic model* is a cause-specific Cox model including sex, age, and comorbidities, while the *Basic model + markers* further included N-terminal pro-brain natriuretic peptide, troponin T, body mass index and resting sinus rate. These analyses were performed in exactly the same population, and due to missing blood biomarker information, the sample comprised a total of 397 patients.

Abbreviations: AF, atrial fibrillation; AUC, area under the receiver operating curve; BMI, body mass index; CABG, coronary artery bypass graft surgery.

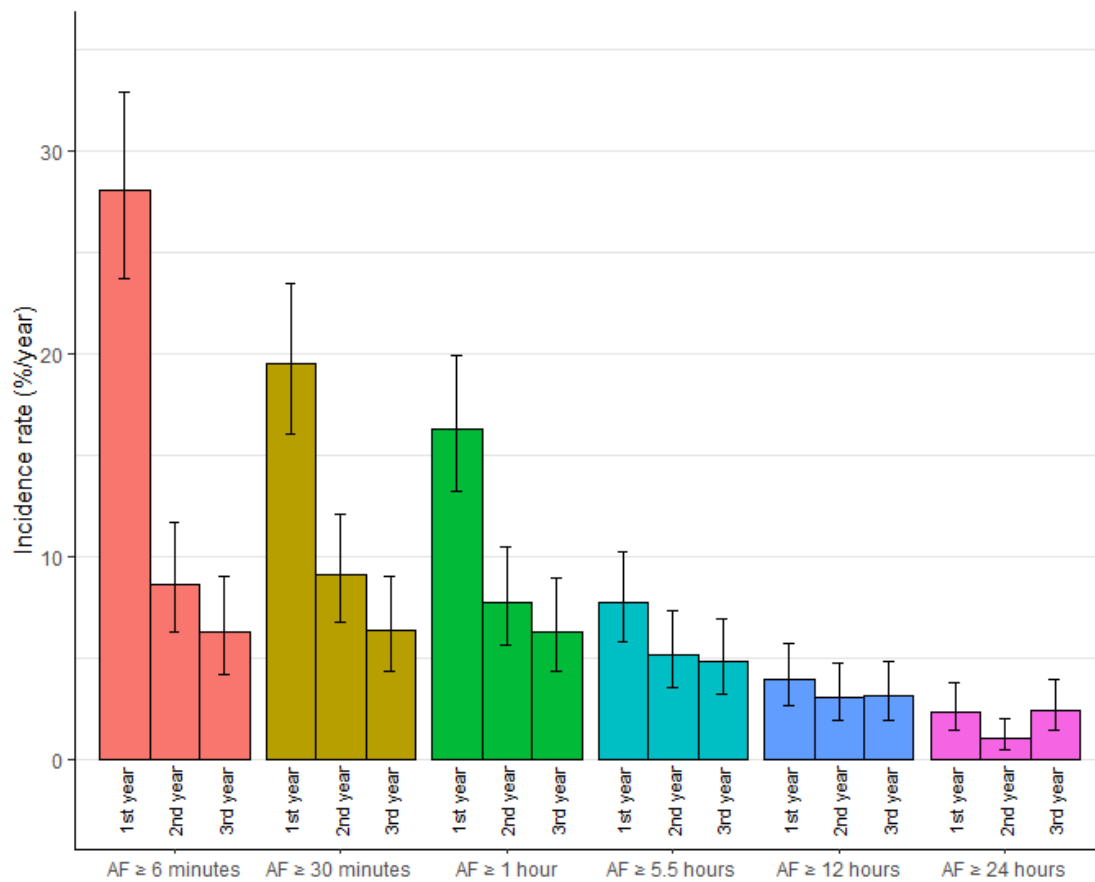
Figure 5 Relationship between baseline markers and incidence rate of AF

Legend: The y-axis presents exact incidence rate (points) and confidence intervals (colored ribbon) for AF episodes lasting ≥ 6 minutes, ≥ 5.5 hours, and ≥ 24 hours in sequentially overlapping subgroups comprising 10% of the population on the x-axis. The curve represents a moving average using local polynomial regression fitting, while the dashed line represents the average incidence rate for the entire population.

Abbreviations: AF, atrial fibrillation; bpm, beats per minute; IR, incidence rate (%/year); NT-proBNP, N-terminal pro-brain natriuretic peptide.

Supplementary Figure 1

Incidence rate of new AF episodes of different durations according to year of monitoring

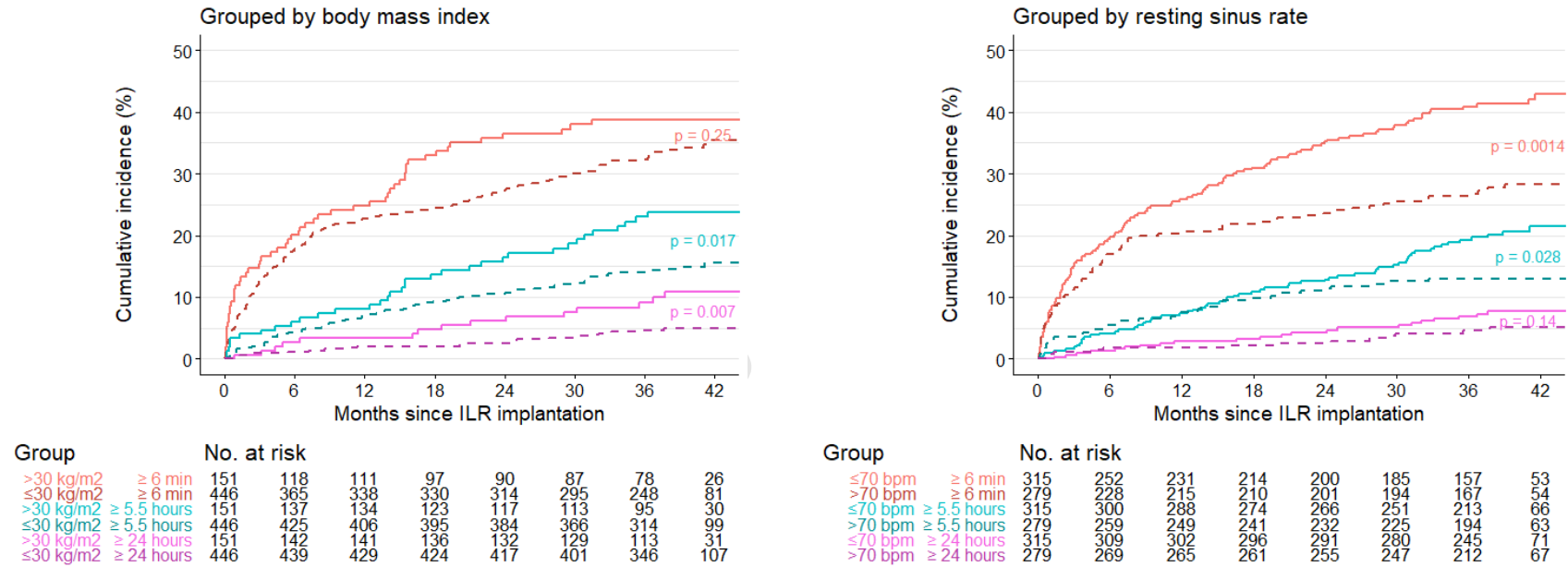


Legend: The figure presents detection rate and exact confidence intervals for new AF episodes of different durations for each of the first three years after device implantation.

Abbreviations: AF, atrial fibrillation.

Supplementary Figure 2a

Cumulative incidence of AF episodes lasting ≥ 6 minutes, 5.5 hours and 24 hours grouped by baseline characteristics

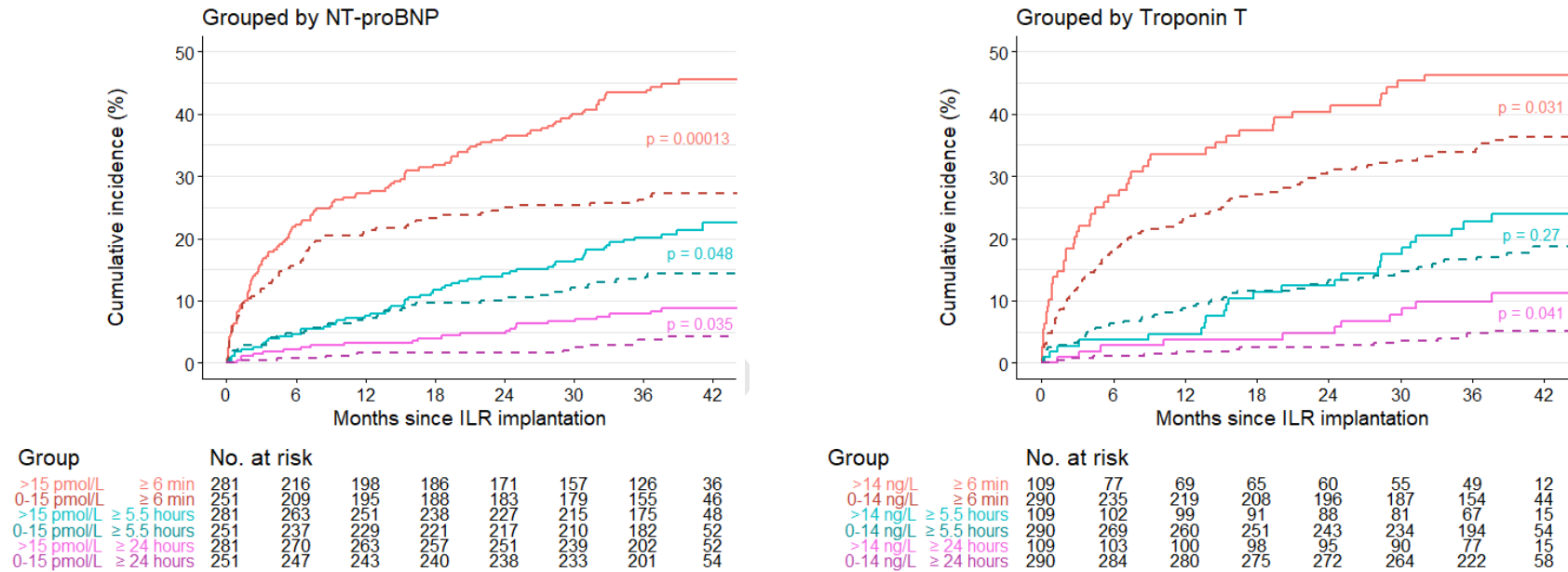


Legend: The figure presents cumulative incidence of AF detection plotted with the Aalen-Johansen method, grouped by BMI and resting sinus rate.

Abbreviations: AF, atrial fibrillation; BMI, body mass index.

Supplementary Figure 2b

Cumulative incidence of AF episodes lasting ≥ 6 minutes, 5.5 hours and 24 hours grouped by baseline characteristics



Legend: The figure presents cumulative incidence of AF detection plotted with the Aalen-Johansen method, grouped by NT-proBNP and troponin T.

Abbreviations: AF, atrial fibrillation; NT-proBNP, N-terminal prohormone of brain natriuretic peptide.

Supplementary Table 1

Baseline characteristics for patients included vs. not included in prediction analyses

Included in prediction analyses n (% of all)	Included 397 (66)	Not included 200 (34)	p
Male sex (%)	235 (59.2)	106 (53.0)	0.18
Age, years (sd)	76.53 (4.40)	75.99 (3.77)	0.14
CHA ₂ DS ₂ VASc score (sd)	3.94 (1.24)	3.89 (1.21)	0.6
CHADS ₂ score (sd)	2.32 (1.06)	2.23 (1.08)	0.3
Heart failure (%)	18 (4.5)	7 (3.5)	0.3
Hypertension (%)	357 (89.9)	183 (91.5)	0.11
Diabetes (%)	120 (30.2)	54 (27.0)	0.7
Previous stroke (%)	73 (18.4)	34 (17.0)	0.6
Previous transient ischemic attack, n (%)	51 (12.8)	16 (8.0)	0.5
Previous systemic embolism, n (%)	25 (6.3)	17 (8.5)	0.8
Previous AMI (%)	46 (11.6)	12 (6.0)	0.10
Previous CABG (%)	21 (5.3)	20 (10.0)	0.4
Valvular heart disease (%)	19 (4.8)	7 (3.5)	0.042
Beta blockers, n (%)	91 (22.9)	53 (26.5)	0.048
Calcium channel blockers, n (%)	129 (32.5)	80 (40.0)	0.6
Nondihydropyridine type, n (%)	9 (2.3)	5 (2.5)	0.4
ACEi, ARB, or renin inhibitors, n (%)	232 (58.4)	124 (62.0)	0.09
Lipid-lowering drugs, n (%)	211 (53.1)	111 (55.5)	1.0
Diuretics, n (%)	118 (29.7)	59 (29.5)	0.5
Platelet inhibitors, n (%)	199 (50.1)	96 (48.0)	0.6
Glucose-lowering drugs, n (%)	102 (25.7)	47 (23.5)	1.0
Systolic BP, mmHg (sd)	151.31 (18.20)	152.70 (19.77)	0.7
Diastolic BP, mmHg (sd)	84.25 (11.31)	86.12 (12.03)	0.6
Resting sinus rate, bpm (sd)	71.14 (11.80)	71.32 (13.58)	0.4
Height, cm (sd)	171.50 (9.20)	169.44 (7.48)	0.06
Body mass index, kg/m ² (sd)	27.68 (4.74)	27.28 (4.16)	0.87
Creatinine, μ mol/L (sd)	88.53 (22.72)	84.35 (25.02)	0.006
NT-proBNP, pmol/L [Q1;Q3]	16.00 [9.00, 28.00]	18.00 [9.00, 32.00]	0.3
hs-CRP, mg/L [Q1;Q3]	2.00 [1.00, 4.00]	2.00 [1.00, 4.00]	0.041
Troponin T, ng/L (sd)	14.58 (6.50)	13.00 (0.00)	0.6

The table presents baseline characteristics of the study participants according to inclusion in prediction analyses (Figure 4 and Supplementary Table 3). Patients could not be included in prediction analyses if they had any missing information about resting sinus rate (3 missing), body mass index (0 missing), NT-proBNP (65 missing), or troponin T (198 missing).

Abbreviations: AF, atrial fibrillation; AMI, acute myocardial infarction; CABG, coronary artery bypass graft surgery; NT-proBNP, N-terminal prohormone of brain natriuretic peptide; hs-CRP, high-sensitivity C-reactive protein.

Supplementary Table 2
Univariate Cox regressions

	AF ≥ 6 minutes	p	AF ≥ 30 minutes	p	AF ≥ 1 hour	p	AF ≥ 5.5 hours	p	AF ≥ 12 hours	p	AF ≥ 24 hours	p
Male sex	1.1 (0.83-1.4)	0.5	1.0 (0.78-1.4)	0.8	1.0 (0.72-1.3)	0.9	1.3 (0.88-2)	0.18	1.7 (0.98-2.9)	0.058	1.6 (0.79-3.1)	0.2
Age (per 5 years)	1.3 (1.1-1.5)	0.0031	1.2 (1-1.4)	0.018	1.3 (1.1-1.5)	0.006	1.2 (1-1.6)	0.046	1.2 (0.87-1.5)	0.3	1.3 (0.9-1.8)	0.2
CHA ₂ DS ₂ VASc >3	1.3 (1-1.8)	0.043	1.2 (0.91-1.7)	0.18	1.4 (1-1.9)	0.044	1.4 (0.94-2.2)	0.099	1.4 (0.8-2.3)	0.3	1.5 (0.74-2.9)	0.3
CHADS ₂ >1	1.6 (1.1-2.3)	0.0068	1.5 (1-2.1)	0.046	1.6 (1.1-2.4)	0.024	1.6 (0.96-2.7)	0.068	1.8 (0.89-3.5)	0.1	1.5 (0.64-3.3)	0.4
Heart failure	0.62 (0.28-1.4)	0.3	0.59 (0.24-1.4)	0.2	0.52 (0.19-1.4)	0.2	0.44 (0.11-1.8)	0.3	0.74 (0.18-3)	0.7	0.62 (0.085-4.5)	0.6
Hypertension	1.2 (0.74-1.9)	0.5	1.2 (0.73-2.1)	0.4	1.2 (0.7-2.2)	0.5	1.1 (0.55-2.2)	0.8	1.2 (0.49-3)	0.68	1.2 (0.38-4)	0.7
Diabetes	1.0 (0.74-1.4)	1.0	0.96 (0.7-1.3)	0.8	1 (0.71-1.4)	1.0	1.1 (0.69-1.6)	0.8	1.0 (0.58-1.7)	1.0	1.0 (0.51-2.1)	0.9
Previous stroke	1.2 (0.88-1.7)	0.2	1.2 (0.85-1.7)	0.3	1.3 (0.9-1.9)	0.16	1.5 (0.92-2.3)	0.11	1.7 (0.98-3.1)	0.057	1.5 (0.73-3.3)	0.3
Previous AMI	1.0 (0.65-1.6)	0.9	1 (0.63-1.7)	0.9	1.1 (0.66-1.8)	0.7	1.4 (0.74-2.5)	0.3	1.5 (0.7-3.1)	0.3	1.5 (0.58-3.9)	0.4
Previous CABG	0.97 (0.56-1.7)	0.9	1.2 (0.68-2)	0.6	1.2 (0.7-2.2)	0.5	1.8 (0.97-3.4)	0.06	2.7 (1.3-5.4)	0.0069	3.5 (1.5-8.1)	0.003
Valvular disease	1.6 (0.9-2.8)	0.11	1 (0.49-2)	1.0	0.82 (0.36-1.8)	0.6	0.85 (0.31-2.3)	0.8	0.73 (0.18-3)	0.6	0.59 (0.081-4.3)	0.6
Beta blockers	0.95 (0.69-1.3)	0.7	1.1 (0.76-1.5)	0.7	1.1 (0.79-1.6)	0.5	1.4 (0.88-2.1)	0.17	2.0 (1.2-3.4)	0.0078	2.2 (1.2-4.3)	0.02
Calcium channel blockers	1.3 (0.99-1.7)	0.06	1.3 (0.94-1.7)	0.11	1.3 (0.91-1.7)	0.17	1.5 (0.99-2.2)	0.055	2.1 (1.3-3.5)	0.0035	1.6 (0.83-3.0)	0.2
Non-dihydropyridine type	2.1 (1.1-4)	0.033	2.2 (1.1-4.6)	0.025	1.9 (0.82-4.2)	0.14	1.4 (0.43-4.3)	0.6	1.4 (0.34-5.7)	0.7	NA	NA
ACEi, ARB, or renin inhibitors	0.97 (0.73-1.3)	0.8	0.96 (0.71-1.3)	0.8	0.96 (0.7-1.3)	0.8	0.9 (0.61-1.3)	0.6	0.91 (0.55-1.5)	0.7	1.3 (0.64-2.5)	0.5
Lipid-lowering drugs	1.3 (0.95-1.6)	0.11	1.3 (0.98-1.8)	0.064	1.4 (1-1.9)	0.037	1.7 (1.1-2.6)	0.011	1.6 (0.93-2.7)	0.088	1.4 (0.74-2.8)	0.3
Diuretics	1.1 (0.85-1.5)	0.4	1.2 (0.84-1.6)	0.4	1.2 (0.88-1.7)	0.2	1.3 (0.85-1.9)	0.2	1.1 (0.63-1.9)	0.8	1.2 (0.58-2.3)	0.7
Platelet inhibitors	1.1 (0.88-1.5)	0.3	1.2 (0.88-1.6)	0.3	1.2 (0.9-1.7)	0.18	1.3 (0.87-1.9)	0.2	1.6 (0.95-2.6)	0.081	1.8 (0.91-3.4)	0.09
Glucose-lowering drugs	0.98 (0.71-1.3)	0.9	1 (0.71-1.4)	1.0	1 (0.71-1.4)	1.0	1.1 (0.72-1.7)	0.6	1.2 (0.66-2)	0.6	1.3 (0.63-2.6)	0.5
Systolic BP (per 5 mmHg)	0.99 (0.92-1.1)	0.7	1.0 (0.93-1.1)	1.0	1.0 (0.93-1.1)	0.8	1.0 (0.9-1.1)	1.0	0.94 (0.82-1.1)	0.3	1 (0.88-1.2)	0.6
Diastolic BP (per 5 mmHg)	0.97 (0.86-1.1)	0.6	0.98 (0.86-1.1)	0.8	0.97 (0.85-1.1)	0.7	0.98 (0.82-1.2)	0.8	0.95 (0.76-1.2)	0.6	1.1 (0.84-1.4)	0.5
Resting sinus rate (per 5 bpm)	0.83 (0.74-0.94)	0.0021	0.83 (0.73-0.94)	0.0028	0.82 (0.71-0.94)	0.0033	0.79 (0.66-0.94)	0.0075	0.81 (0.65-1)	0.06	0.74 (0.55-0.99)	0.04

Height (per 10 cm)	1.1 (0.93-1.3)	0.3	1.1 (0.89-1.3)	0.5	1.0 (0.85-1.2)	0.8	1.1 (0.89-1.4)	0.4	1.3 (0.99-1.8)	0.06	1.3 (0.9-1.9)	0.2
Body mass index (per 5 kg/m ²)	1.1 (0.96-1.3)	0.17	1 (0.89-1.2)	0.7	1.1 (0.93-1.3)	0.3	1.2 (0.98-1.5)	0.08	1.2 (0.94-1.6)	0.13	1.5 (1.1-2.0)	0.01
Creatinine (per 20 µmol/L)	1.0 (0.94-1.2)	0.4	1.1 (0.94-1.2)	0.3	1.0 (0.92-1.2)	0.5	1.1 (0.95-1.3)	0.2	1.2 (0.99-1.4)	0.06	1.2 (0.93-1.4)	0.2
NT-proBNP (per doubling)	1.2 (1.1-1.3)	0.0001 8	1.2 (1.1-1.4)	0.0002 8	1.2 (1.1-1.4)	0.0009 3	1.2 (1.1-1.4)	0.004 3	1.4 (1.2-1.6)	0.0003 5	1.4 (1.1-1.8)	0.002
hs-CRP (per doubling)	1.1 (1-1.3)	0.039	1.1 (0.94-1.2)	0.3	1.1 (0.95-1.3)	0.2	1.1 (0.91-1.3)	0.4	1.0 (0.8-1.3)	0.8	1.2 (0.88-1.6)	0.3
Troponin T (per 10, ng/L)	1.4 (1.1-1.7)	0.0006 8	1.3 (1.1-1.6)	0.0067	1.4 (1.1-1.7)	0.0042	1.5 (1.1-1.9)	0.004 5	1.6 (1.2-2.2)	0.0007 8	1.9 (1.4-2.7)	0.000 2

The table presents hazard ratios (with 95% confidence intervals) and p-values from univariate Cox regressions with time to AF of different durations as endpoint.

Abbreviations: AF, atrial fibrillation; AMI, acute myocardial infarction; BP, blood pressure; CABG, coronary artery bypass graft surgery; NT-proBNP, N-terminal prohormone of brain natriuretic peptide; hs-CRP, high-sensitivity C-reactive protein; NA, not applicable.

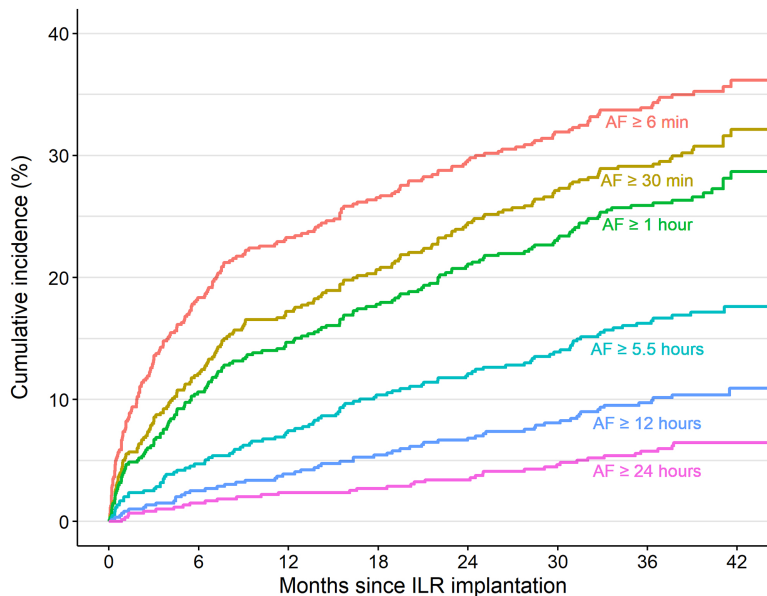
Supplementary Table 3

AUC for prediction of AF after split in training and test sets

Endpoint	Model	AUC	Δ AUC	P
AF \geq 6 min	Basic model	0.594 [0.45;0.74]	0.045 [-0.06;0.15]	0.4
AF \geq 6 min	Basic model + markers	0.638 [0.49;0.79]		
AF \geq 5.5 hours	Basic model	0.569 [0.40;0.74]	0.056 [-0.07;0.17]	0.3
AF \geq 5.5 hours	Basic model + markers	0.614 [0.44;0.79]		
AF \geq 24 hours	Basic model	0.499 [0.18;0.82]	0.221 [0.08;0.36]	0.002
AF \geq 24 hours	Basic model + markers	0.719 [0.46;0.98]		

The table presents time-dependent area under the receiver operating characteristics curve (AUC) for two models in prediction of atrial fibrillation applied in the same test dataset, along with differences between AUC (Δ AUC) and P-values for difference. The *Basic model* is a cause-specific Cox model sex, age, and comorbidities, while the *Basic model + markers* further included N-terminal pro-brain natriuretic peptide, troponin T, body mass index and resting sinus rate.

Abbreviations: AF, atrial fibrillation; AUC, area under the time-dependent receiver operating curve; CABG, coronary artery bypass graft surgery.



AF duration No. at risk

≥ 6 min	597	483	449	427	404	382	326	107
≥ 30 min	597	519	484	460	432	408	348	110
≥ 1 hour	597	528	498	476	451	430	363	115
≥ 5.5 hours	597	562	540	518	501	479	409	129
≥ 12 hours	597	575	561	544	529	510	437	135
≥ 24 hours	597	581	570	560	549	530	459	138

Figure 1

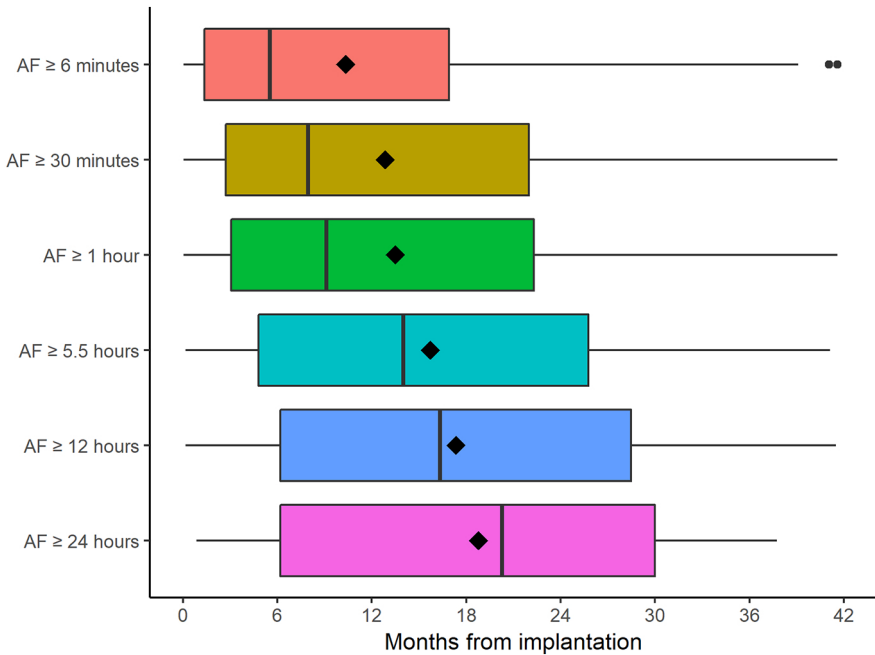


Figure 2

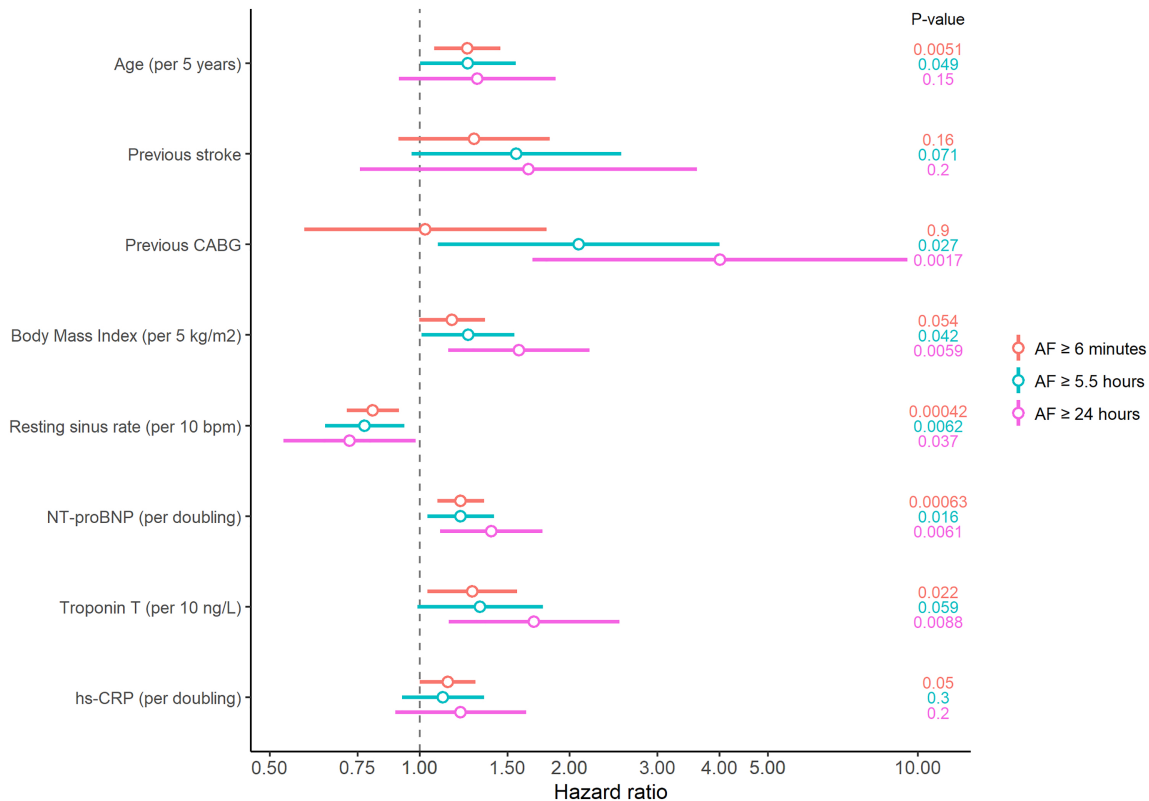
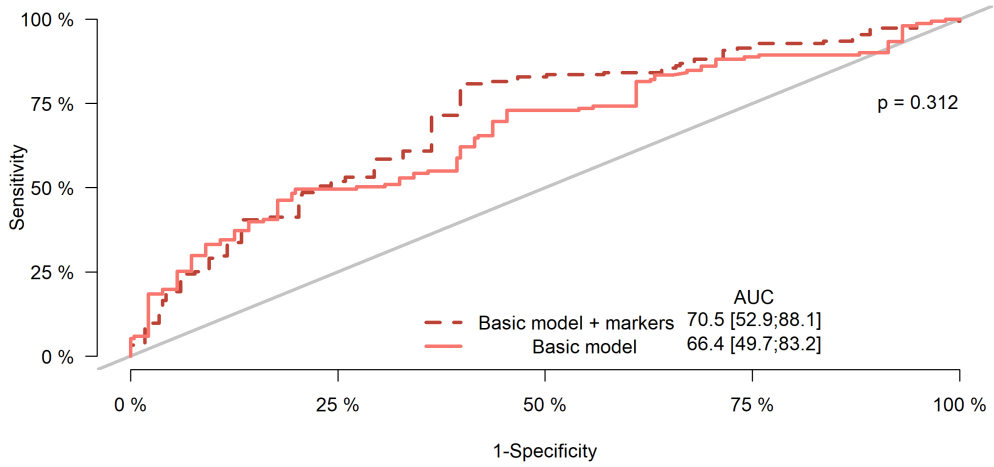
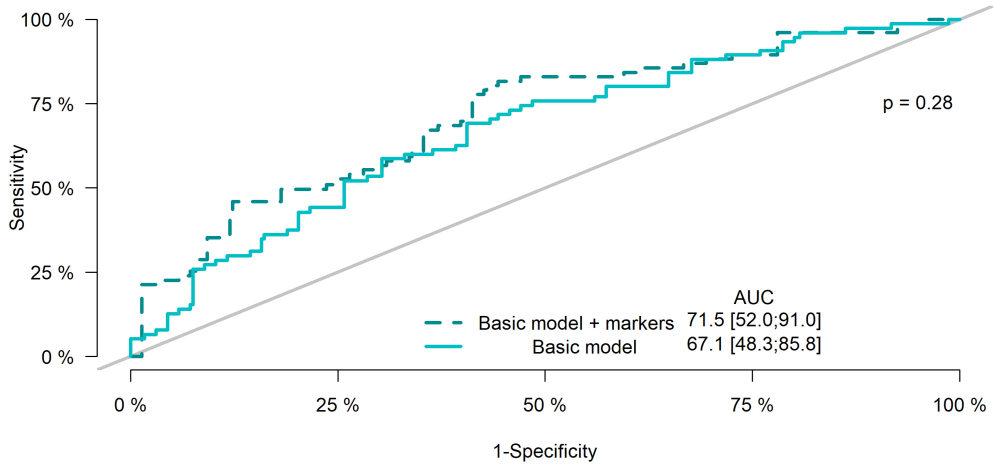


Figure 3

Prediction of AF ≥ 6 min



Prediction of AF ≥ 5.5 hours



Prediction of AF ≥ 24 hours

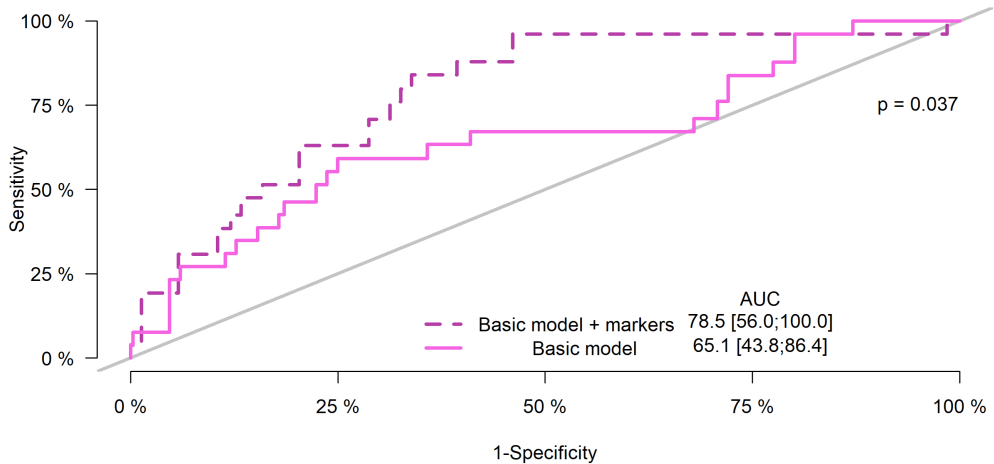


Figure 4

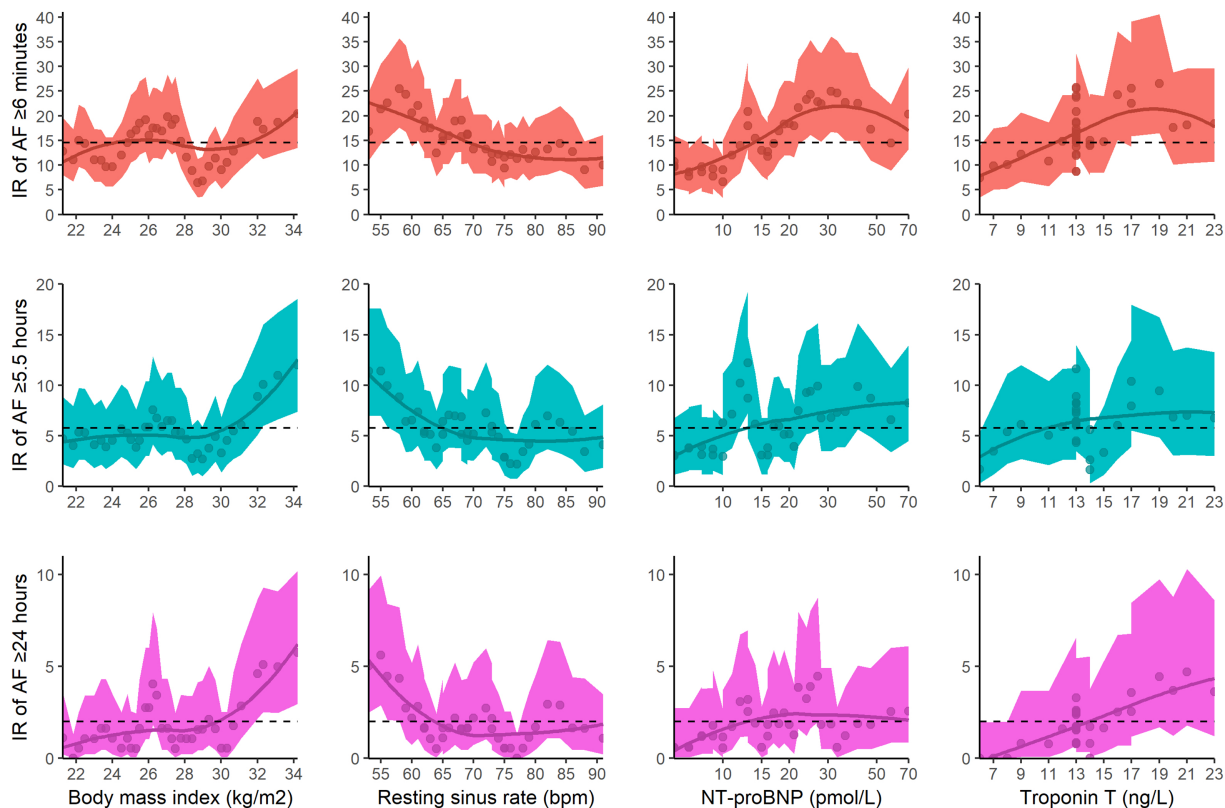


Figure 5